1 ORIGINAL ARTICLE

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- 3 Association between woman authorship and woman editorship in infectious diseases
- 4 journals: a cross-sectional study
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23 **Research in context**

24 Evidence before this study

Gender disparities are still pervasive in academic medicine. Among the many 25 domains affected, the differential publication output of women is one of the best-26 documented aspects across medical disciplines. Recently, the underrepresentation 27 of women in editorial boards has come to renewed attention emphasizing the 28 29 powerful and strategic role editors embody. To our knowledge, the association between women editorship and women authorship has been under-investigated in 30 infectious diseases thus far. 31 32 We searched PubMed/Medline for articles using the keywords "gender AND (equit" OR disparit* OR inequit* OR parit* OR gap) AND (academic OR publish* OR 33

34 publication* OR authorship OR editor*) AND (medicine OR medical)" until May 20,

35 2022, without language restrictions.

36 Added value of this study

We investigate the complex interrelations between first and last authorship each with 37 editorship, while taking a journal's impact factor into account. In our analyses, we 38 39 show a positive association between woman editorship and woman first and last 40 authorship in infectious diseases journals. Interestingly, the woman-to-man ratio in first authors in infectious diseases journals is nearly equal. In contrast, woman last 41 authors in infectious disease journals are underrepresented. Surprisingly, the 42 43 proportion of woman editors in infectious diseases journals was comparable to obstetrics/gynaecology journals, despite the higher proportion of women clinicians in 44 45 obstetrics/gynaecology.

46 Implications of all the available evidence

Differences in woman editorship partially explained gender disparities in first and last
authorship in this study. Our data may support the notion that woman editorship

- 49 potentially serves as a possible lever to increase woman authorship proportions in
- 50 academic medicine journals.

51 ABSTRACT

52 Background:

53 Gender inequity is still pervasive in academic medicine, including publishing in

54 scientific journals. We aimed to (i) ascertain the proportion of women among first and

55 last authors and editors in infectious diseases (ID) journals and (ii) assess the

56 association between woman editors and both woman first and last authors while

57 controlling for a journal's impact factor (IF).

58 Methods:

59 In this observational cross-sectional study, we randomly selected 40 ID journals (ten

60 from each IF quartile) and 20 journals (five from each IF quartile) in each

61 "obstetrics/gynaecology" (OB/GYN) and "cardiac/cardiovascular systems" (CARDIO)

62 as a comparator. First and last authors' names of all citable articles published during

63 2018/2019, and names of each journals' editors (including editors-in-chief, EiC) were

retrieved; genders were predicted with genderize.io.

65 Findings:

A total of 11,027 ID articles were analysed, yielding a women-to-men ratio of 49.3% 66 (5,350/10,853) vs 50.7% (5,503/10,853) among first authors (first author gender 67 68 indeterminate in 1.5% [167/11.027] of the cases), and 34.9% (3.788/10.865) vs 65.1% (7,077/10,865) among last authors (last author gender indeterminate in 1.4% 69 70 [155/11,027] of the cases), whereas seven articles had no author indexed. Of 495 ID journal editors, 32.3% (160/495) were women, and 37.5% (15/40) of EiC were 71 72 women. Quasi-Poisson regression estimated a significant effect of woman editors on 73 woman last authors (incidence rate ratio, IRR, 1.92; 95% CI 1.45–2.55; p<0.001) and on woman first authors (IRR, 1.32; 95% CI 1.06–1.63; p=0.012) in ID journals. The 74 journal's IF exerted no effect in these analyses. 75

76 Interpretation:

83	Key words:
82	Funding: European Society of Clinical Microbiology and Infectious Diseases.
81	representation.
80	medicine and suggest a need for revised policies towards increased woman editorial
79	explain the gender disparities observed in women publishing activities in academic
78	last and first authors in the analysed ID journals. These findings may contribute to
77	The proportion of woman editors appears to influence the proportion of both woman

- 84 Gender disparity, gender inequity, gender inequality, academic medicine, publishing,
- 85 authorship, editorship, infectious diseases

86 **INTRODUCTION**

87 It is a sobering reality that women still suffer from gender disparity within academic medicine despite representing more than half of the health workforce.¹ Women 88 publish less than men^{2,3} and in less influential journals,^{4,5} are less likely to be 89 subsequently cited⁶ and to hold an influential editorial position in medical journals.⁷ 90 There are many more domains where underrepresentation of women is evident, e.g., 91 disparities among conference speakers and chairs, and guideline group leads,^{8,9} 92 leading to self-perpetuating hurdles that aggravate these disparities.¹⁰ Ultimately, 93 women have been shown to be less likely to reach seniority in academic positions, 94 95 *i.e.*, full professorship, than their man counterparts, also in the field of infectious diseases (ID).^{3,11} 96

97 As a surrogate, publishing output exemplifies academic productivity and career advancement like no other metric.^{2,12} A high publication output is critical for 98 promotions to professorship and leadership positions, and for receiving grants, 99 100 awards, and speaker invitations.¹³ Vice versa, progression on the ladder of academic 101 hierarchy is crucial for further research output, possibilities for collaborations, and visibility in the scientific community. Of note, hierarchy levels within the publishing 102 system also influence academic advancement,¹² rendering senior research positions 103 even more susceptible to power-based gender disparities. 104

105 Gender disparity is not only found in authorship. Women continue to be

106 underrepresented among editorial positions of medical journals and especially in ID

journals, despite a continuous increase during the past decades.^{7,14} Furthermore,

108 women are underrepresented during the peer-review process as they are less likely

to be invited to review articles.¹⁵ Scientific gatekeeping by men in their strategic

110 decision-making roles as editors and reviewers may contribute to the finding that

111 women are less likely to publish in high-impact journals.^{5,7} The gender of a journal's

editors and/or chief editor may partly influence how likely a manuscript is considered 112 113 for peer-review and publication. Factors which may directly or indirectly influence the editors' decisions include possible same-gender preference by editors,¹⁶ as well as 114 115 differences between man and woman authors regarding rhetorical style, selfassuredness of reportage, and preference for certain research methods or topics.¹⁷ 116 Overall, the complex relationship between women underrepresentation among 117 118 editors as well as gender disparity among authors remains disputed, with causal mechanisms at play between those two domains not yet fully untangled.¹⁸ There is a 119 120 strong need within the scientific community to diagnose the underlying causes of these gender-based disparities in order to identify barriers and possible mechanisms 121 122 and to alleviate gender gaps in publication output and representation in research. In our study focusing on ID journals, we hypothesized that a high proportion of 123 124 woman editors would be associated with a high proportion of woman first authors and last authors. We aimed to (i) ascertain the proportion of women among first and last 125 authors and editors in ID journals and (ii) assess the association between woman 126 127 editors and both woman first and last authors while controlling for a journal's impact factor (IF). 128

129 **METHODS**

130 Study design and data sources

Due to the lack of a basis for meaningful sample-size calculations, we assumed that 131 132 an association would be detectable by selecting at least 40 out of the 93 journals publishing human clinical ID research data with an IF indexed in the Journal Citation 133 134 Reports (Clarivate Analytics) category of "Infectious Diseases" (ID). To confer an 135 even distribution among IF quartiles, we randomly selected ten journals from each 136 2020 IF guartile as provided by the Journal Citation Reports and Web of Science (both Clarivate Analytics), using an R code to generate random numbers. Journals 137 138 were included when they had an IF and when all first and last author given names were retrievable. Journals were excluded when the first and last author given names 139 140 were systematically not provided via Clarivate and/or when they had only one editor 141 listed. As comparator groups, we chose obstetrics/gynaecology (OB/GYN) as a speciality with a known high share of women workforce, and cardiology ('cardiac and 142 143 cardiovascular systems' [CARDIO]), a specialty with a low women workforce 144 (appendix p 3). For both comparator groups, we selected 20 journals each (out of 83 OB/GYN journals and 142 CARDIO journals), five from every IF quartile. We then 145 146 retrieved all citable articles published during 2018 and 2019 for each of these journals via Journal Citation Reports and Clarivate Web of Science (Clarivate, 2021, 147 all rights reserved). All citable articles refer to the total number of articles retrievable 148 via Clarivate Web of Science for the years 2018 and 2019 that were counted towards 149 150 the 2020 impact factor.

We then extracted all given and family names of the first and last authors of every article in downloading the excel files and manually curating the data. We utilized the genderize.io interface for Google Sheets and a self-programmed Python algorithm to estimate gender of each first and last author. Genderize.io is a tool utilized for binary gender prediction in previous studies with high accuracy rates for both man and
woman first names.¹⁹ In addition, we collected names of all editors-in-chief (EiC),
deputy editors, section editors, and associate editors, for the years 2018 and 2019,
and performed gender determination as described above (appendix p 3). The
reporting of this study adhered to the STROBE statement (appendix pp 6-7).

160 Data quality and integrity

161 Data extraction was performed by two members of the study team in parallel, who performed cross-checks on 20% of each other's collected data. When genderize io 162 could not determine a gender, two members of the study team independently 163 164 performed an additional internet search (appendix, p 3). To account for potential inaccuracies of genderize io for predicted gender probabilities below 80%, we 165 166 performed a subgroup analysis ("conservative cohort") with all articles for which a 167 gender probability of \geq 80% could be attained. For this conservative cohort, all genders predicted with a probability at the cut-off of 80% were additionally searched 168 169 as described above.

170 <u>Outcomes</u>

171 The variables of interest were woman first authors and woman last authors. Factors 172 with potential influence on the gender of authors included woman editor proportion. woman EiC status, and journal IF, all of which were included in the analyses as 173 control variables. Additionally, we included the variable women workforce share for 174 175 supplementary analyses which was operationalized as percentage of woman 176 clinicians in the workforce of selected countries (appendix p 3). We calculated the 177 proportion of woman first authors as the number of articles with a woman first author 178 divided by the sum of articles with both woman and man first authors in a given journal. Articles without any authors or cases in which gender could not be 179 determined were not included in this calculation. 180

Likewise, the proportion of woman editors was calculated as the number of woman editors (including EiC, deputy editors, section editors, associate editors) divided by the sum of both woman and man editors in a given journal. The variable woman EiC was determined as a binary outcome per journal.

185 Statistical analysis

186 Differences in the number of woman EiC (including Co-editors-in-chief) between

187 high-impact journals and low-impact journals were compared using Fisher's exact

test. We used Spearman rank correlation to test associations between woman

editorship and woman authorship as well as between journal IF and woman

authorship. We assessed the association between woman editorship and woman first

and last authorship by fitting a quasi-Poisson regression model and including IF as a

192 control variable. Logarithmic transformation was used for non-normally distributed

193 predictor variables. R Studio (Version 1.3.1093), Python (3.9.6), and Prism 9.0.0

194 (GraphPad) were used for statistical analyses. The statistical significance level was195 set at 0.05.

195 Set at 0 05.

196 Role of the funding source

197 The funding source had no role in study design, data collection, data analysis, data 198 interpretation, or writing of the report. The corresponding author had full access to all 199 the data in the study and had final responsibility for the decision to submit for 200 publication.

201 **RESULTS**

- 202 We analysed 11,027 articles from 40 indexed ID journals alongside 6,450 and 7,157
- articles from 20 OB/GYN and 20 CARDIO journals, respectively (table 1).
- Of the 11,027 ID articles, 167 (1.5%) had an indeterminable first author gender, and
- 155 (1.4%) had an indeterminable last author gender, and 7 (0.1%) no authors
- indexed (Table 2). Among first authors, 5,350 (49.3%) were predicted to be women
- and 5,503 (50.7%) were men. Among last authors, women and men accounted for
- 208 3,788 (34·9%) and 7,077 (65·1%), respectively.
- 209 Pertaining to ID journals, we identified 577 editors; of these, 190 (32.9%) were
- identified as women and 387 (67.1%) as men (table 1).
- In comparison, the proportion of woman first authors was higher in OB/GYN journals
- 212 (58.2%; 3,696/6,347; first author gender indeterminate in 51 cases; no authors
- indexed in 52 cases) and lower in CARDIO journals (28.6%; 2,015/7,056; first author
- gender indeterminate in 86 cases; no authors indexed in 15 cases), while the
- proportion of woman last authors were 43.6% (2,924/6,303; last author gender
- indeterminate in 95 cases) in OB/GYN journals and 16.9% (1,186/7,027; last author
- 217 gender indeterminate in 115 cases) in CARDIO journals (table 1). Woman editor
- proportions were similar in OB/GYN journals (34.1%; 103/302), and lower in CARDIO
- 219 journals (16.8%; 93/553), compared to ID (32.9%, 190/577).
- 220 When comparing the number of woman EiCs between specialties, there was a non-
- significantly lower number of woman EiCs in ID journals than OB/GYN (13/40 vs
- 12/20; p=0.055), and a significantly higher number compared to CARDIO journals
- 223 (0/20; p=0.003) (table 1; appendix p 5).
- In a first step, a bivariate association was estimated to approximate crude effects.
- 225 There was a significant positive association between the proportion of woman editors
- and a) woman first authors ($R_s 0.62$, 95% CI 0.38-0.79; p<0.0001) and b) woman

last authors in ID journals ($R_s 0.63$, 95% CI 0.38-0.79; p<0.0001) (figure 1).

Additional bivariate associations are shown in the appendix (p 4).

For the aggregated dataset, *i.e.*, combined data of ID, OB/GYN, and CARDIO, we 229 230 observed an association between woman last author proportion and woman first author proportion (figure 2). 231 232 We performed a guasi-Poisson regression analysis on an aggregated model for all 233 three specialties (ID, OB/GYN, CARDIO). We found that for the dependent variable 234 woman first authorship, the factor with the strongest impact was the proportion of woman editors (incidence rate ratio, IRR 1.80, 95% CI 1.32 - 2.44, p<0.001), 235 236 alongside woman last authorship (IRR 1.26, 95% CI 1.09 – 1.45, p=0.002) and woman EiC (IRR 1.22, 95% CI 1.08 – 1.39, p=0.002). Journal IF was not significantly 237 associated with woman first authorship (IRR 0.92, 95% CI 0.78 – 1.09, p=0.346). 238 239 Woman last authorship as dependent variable was significantly associated with

240 woman editorship (IRR 2.56, 95% CI 1.70 – 3.85, p<0.001), woman EiC (IRR 1.29,

241 95% CI 1.09 – 1.52, p=0.003), and woman first authorship (IRR 1.28, 95% CI 1.05 –

1.58, p=0.017). Again, journal IF showed no significant association (IRR 0.90, 95%)

243 Cl 0.72 – 1.14, p=0.395) (table 3).

Quasi-Poisson regression analyses specifically for ID journals yielded that woman first authorship as dependant variable was significantly associated with the proportion of woman editors (IRR 1.32, CI 1.06 - 1.63, p=0.012), while women EiC, woman last authorship or journal IF had no significant effect (table 4).

248 When woman last author was used as a dependent variable, we observed again that 249 the proportion of woman editors was significantly associated with woman last

authorship (IRR 1.92, 95%CI 1.45–2.55, p<0.001), with again woman EiC, journal IF,

and woman first author reaching no significant effect (table 4).

252 Supplementary regression analyses included the control variable woman "workforce share", which was a significant predictor of the proportion of woman in first and last 253 254 authorship. Of note, woman editorship was also a significant predictor of both woman first and last authorship in these analyses, independently from the woman workforce 255 256 share (appendix p 8, table S2). 257 Regression analyses on the conservative ID cohort (*i.e.* predictive gender probability above 80%) yielded comparable results (appendix p 8, table S3) compared to the 258 259 main analyses. Results of the specific regression analyses regarding the OB/GYN 260 and CARDIO journals are shown in the appendix (p 9), alongside the journal

characteristics (p 10).

262 **DISCUSSION**

263 We found that the woman-to-man ratio in first authorship in 40 selected ID journals was nearly equal (49.3% vs 50.7%), a laudable finding which demonstrates near 264 gender equity for the years under consideration. As this partially contradicts previous 265 research regarding gender differences in first authorship in the field of ID,³ this 266 differential finding may be due to reasons such as sample size differences or 267 268 sampling of the journals, also exemplified by the large range of woman first author 269 proportions observed in our dataset. An additional factor for explanation may have 270 been recent efforts by ID journals and societies to increase women representation at all levels of the publication submission and review process.^{11,20} 271 272 In contrast, the woman-to-man ratio in last authorship was considerably lower in ID 273 journals (34.9% vs 65.1%) in our analyses, which is in line with previous research in other medical specialties.⁴ More rigorous efforts seem to be warranted to increase 274 275 research output by women in senior author positions, since first authorship appears 276 to not automatically translate into last authorship later in the career ("senior author gender gap")⁴ due to, e.g., gender-based task specialization and contributorship in 277 research teams, besides other reasons.²¹ In general, last authorship necessitates 278 279 more seniority in research experience, international recognition, and expertise, which accumulates over the course of a career, but depends on additional factors within the 280 publication process and academic medicine's structural setup as well.²² 281 Furthermore, we found that the proportion of woman editors in ID journals (32.9%)282 283 was similar to OB/GYN (34.1%) and higher as in CARDIO journals (16.8%). Overall, gender parity among editors clearly remains elusive in all three journal categories. 284 The nearly similar proportion of woman editors in ID as in OB/GYN journals is a 285 surprising finding given the substantially higher proportion of women clinicians 286 working in OB/GYN and publishing as first and last authors. For ID and academic 287

288 medicine in general, unequal gender representation in editorial positions has been
 289 demonstrated before.^{3,14,23} Our results complement this finding.

In our regression models, woman editorship was significantly associated with woman 290 291 first and last authorship: the higher the proportion of woman editors across ID journals, the higher the proportion of woman first and last authors in the analysed 292 293 time period. This effect was also observed in the aggregated data model across all 294 ID, OB/GYN, and CARDIO journals, albeit additional factors showing significance in 295 this model as well, *i.e.*, EiC. Journal IF was not associated with first or last authorship 296 in any of the ID or the aggregated model which contradicts previous reports on first authorship, but complements data on last authorship.⁴ 297

The reciprocal association between woman first and last author observed in the 298 299 aggregated model is suggestive of endogeneity that we did not analyse further, and 300 which would necessitate a more complex statistical approach. Likewise, as our supplementary analyses show, the impact of the proportion of women in the 301 302 healthcare workforce of ID, OB/GYN and CARDIO on authorship is unsurprising. 303 However, woman editorship was still independently predictive for woman authorship despite including workforce share as control variable in the model. 304 305 Overall, our findings indicate a potential effect of woman editorship on woman 306 authorship. Among other reasons, one possible explanation for this is unconscious and implicit gender bias among editors when judging a submitted article.²⁴ This has 307 308 been demonstrated in a recent study within gastroenterology, in which editorial 309 boards dominated by men were correlated positively with man first and last authorship.¹⁶ The same seem to apply to grant funding assessments,²⁵ in which 310 311 women were rated less favourably than their man counterparts despite similar intellectual content. On the contrary, in a pilot study on one ID journal, articles 312

submitted by women were more likely to be accepted by both woman and man

editors than those submitted by man authors.²⁶ A more recent analysis across
several disciplines, including biomedicine and health, seems to support this finding.¹⁸
However, both studies have limitations such as excluding desk-rejections as unit of
analysis.

Certainly, other factors in the peer-review process may be at play and may aggravate this gender bias. This includes differences between men and women authors regarding rhetorical style, self-assuredness of reportage, or questioning editorial decisions of rejection.^{17,27} Furthermore, the editors' and reviewers' preference for authors from prestigious institutions or with previous reputation in the field under consideration as well as preference for certain research methods may create implicit and indirect gender bias.²⁸

Many decisions in the publishing and especially the peer-review process are not yet 325 transparent enough to prevent all-men and authoritative circles to dominate and 326 influence decision-making.⁵ In competitive research and academic settings, many 327 328 roles, gender norms, and rules are often aligned with bold and self-assertive behaviour, disproportionately privileging men.^{19,27} Ultimately, woman authors may 329 330 feel discouraged to submit their research work to an all-men editorial board or to a man EiC; and more senior woman researchers may refrain from applying to man-331 dominated editorial positions or a man EiC. 332

Our study has strengths which contribute to the robustness of the findings. Firstly, we included a large share of indexed ID journals, *i.e.*, 40 out of 93 with an IF indexed in the Journal Citation Reports, thereby reaching meaningful statistical results with a high point estimate and minimizing the risk of chance findings. Secondly, we aimed to represent all journal ranks by including ten journals per IF quartile. Furthermore, we enriched the gender prediction process with manual screening, which contributed to the overall very low rate of indeterminate author names. In addition, we conducted

supplementary analyses on a subgroup with very high genderize io prediction 340 341 probability (*i.e.*, \geq 80%). These results were comparable to our main results. There are limitations to our study. One limitation is the lack of information on the 342 343 number and gender distribution of submitted manuscripts as a denominator for the published papers. Furthermore, in a minor proportion of authors, gender could be 344 345 estimated neither by genderize io nor by individual internet research, especially in 346 journals with a high share of authors with gender-neutral names. However, this 347 indeterminate rate was ultimately very low (<2%), due to our aforementioned approach. Our methods were unable to account especially for self-perceived gender 348 349 identities that diverge from the binary gender their first names suggest. This individual gender conceptualization would only have been ascertainable with an 350 351 accompanying survey complementing the genderize io analysis. Given the large 352 dataset, individual author consultation was not manageable, and we surmise that resulting inaccuracies are overall rather minor. Another limitation is that we analysed 353 354 all citable articles without sub-categorising, e.g., in original research articles or 355 reviews. Similarly, we did not analyse time trends, an endeavour beyond the scope of 356 this work, but desirable to conduct in the future, nevertheless. Lastly, we did not take 357 into account the country of origin of authors; the number of total authors and their 358 gender composition; and/or corresponding author gender, all of which may have independent effects on publishing outcome but also influence the gender of the first 359 and last author.27 360

Gender parity in academic medicine remains a yet to be achieved multi-stage ideal
and a challenge at the same time. Our findings show near gender parity in first
authorship in ID journals, while equal gender representation is far from achieved
regarding last authorship and editorship.

This partial success should be lauded, but also framed as call for action to extend 365 366 gender equity to the remaining domains such as last authorship and editorial representation. Sustained efforts are needed over time to translate gender equity in 367 first authorship into increased woman last authorship representation, especially in 368 light of the setbacks for woman researchers' output caused by the Coronavirus 369 disease 2019 pandemic.¹⁹ Efforts to increase the number of women among EiCs and 370 371 editors may positively influence gender representation among first and last authors, 372 as indicated by our aggregated analyses across all three journal categories. Among the actions taken within the past years, the composition of editorial teams 373 374 may have been undervalued as potential mechanism to increase woman authorship. Likewise, increasing the number of woman EiCs may serve as a valuable tool to 375 increase the overall proportion of woman editors.¹⁴ Based on our analyses, gender 376 377 balance among editors and EiC positions may be a promising policy tool to help 378 counteract disparity in academic publishing especially regarding last authorship 379 publications.

380 The effects of such numeric gender parity may extend well beyond the mere concept of checks and balances. In analogy to the observations made during scientific 381 conferences,⁹ it may create visibility of women and generate political momentum, 382 383 potentially contributing to gender equity on other levels as well.²⁹ Gender disparities at all levels in medicine also have profound implications for society, as diversity in 384 leadership positions is needed to reach balanced decisions and produce more 385 representative health science.¹⁹ Gender equity is therefore, besides being a matter of 386 justice and rights for all, decisive for diversifying science to produce more rigorous 387 388 research³⁰ and provide the best possible care to patients. Furthermore, gender inequities in academic publishing should not be framed as a personal shortcoming at 389 390 an individual level, but as a systemic problem affecting half of medicine's workforce.

- 391 Further quantitative and qualitative studies are needed to better illuminate the
- 392 intrinsic mechanisms and modalities of editorial decision-making processes,
- specially with regards to the share and success of submissions authored by women.

394 Contributors

- 395 KL and CP conceived of and conceptualized the study. KL, LH, SEM, and CP
- collected the data. KL and CP accessed and verified all the data. KL, CP and MW
- 397 conducted the statistical analyses. All authors interpreted the data. KL and CP wrote
- the first draft of the manuscript. All authors critically revised the manuscript and
- approved the final version for submission.
- 400 Data sharing
- 401 Data obtained during this study are subject to copyright regulations and thus cannot
- 402 be shared.

403 **Declaration of interests**

404 We declare that we have no conflicts of interest.

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TABLES

489 490 **Table 1.** Number of screened articles, journals, respective impact factors, and gender distribution among first authors, last authors, and editors, for journals from the different fields of infectious diseases (ID), obstetrics and gynaecology (OB/GYN), and cardiac and cardiovascular systems (CARDIO); IQR: interquartile range.

	ID	OB/GYN	CARDIO
Citable articles from 2018 and 2019 analysed	11,027	6,450	7,157
Number of journals analysed	40	20	20
Median impact factor (IQR)	3·625 (2·632– 5·571)	2.839 (2.004–3.981)	3·067 (2·550– 5·550)
Articles with woman first author, total	5,350	3,696	2,015
Articles with man first author, total	5,503	2,651	5,041
Articles with indeterminate first author gender, total	167	51	86
Articles without author, total	7	52	15
Woman:man first author ratio	49·3% : 50·7%	58·2% : 41·8%	28.6% : 71.4%
Articles with woman last author, total	3,788	2,747	1,186
Articles with man last author, total	7,077	3,556	5,841
Articles with indeterminate last author, total	155	95	115
Woman:man last author ratio	34.9% : 65.1%	43.6% : 56.4%	16·9% : 93·1%
Number of woman editors, total	190	103	93
Number of man editors, total	387	199	460
Number of woman chief editors (%)	13 (32·5%)	12 (60.0%)	0 (0.0%)
Woman:man editor ratio	32·9% : 77·1%	34·1% : 65·9%	16.8% : 83.2%

- 494 495 **Table 2.** Characteristics of 40 ID journals analysed within this study. Number of first authors with indeterminate gender: 167; number of last authors with indeterminate gender: 167; number of articles without author: 10. EiC: editor-in-chief; IF: impact factor.

	2020 IF	citable articles	first authors woman	first authors man	First authors woman proportion	Last authors woman	Last authors man	Last authors woman proportion	editors woman	editors man	editors woman proportion	EiC woman
Lancet ID	25.071	283	119	164	42.0%	88	195	31.1%	1	4	20.0%	no
Lancet HIV	12.767	133	57	76	42.9%	55	78	41.4%	2	2	50·0%	no
Clin Infect Dis	9.079	1,146	531	609	46.6%	381	753	33.6%	7	10	41.2%	no
J Travel Med	8.490	102	53	48	52.5%	22	78	22.0%	12	24	33.3%	yes
Emerg Infect Dis	6.883	667	325	342	48.7%	206	451	31.4%	5	26	16.1%	no
Eurosurveillance	6.307	384	219	164	57·2%	175	205	46·1%	15	11	57.7%	yes
Travel Med Infect Dis	6.211	133	60	71	45.8%	42	89	32.1%	3	6	33.3%	yes
J Infect	6.072	237	103	126	45·0%	73	161	31.2%	3	12	20.0%	no
Int J Hyg Envir Heal	5.840	243	136	103	56·9%	94	146	39.2%	6	7	46.2%	yes
J Antimicrob Chemother	5.790	980	474	482	49.6%	305	650	31.9%	2	9	18.2%	no
Curr Opin Infect Dis	4.915	165	75	90	45·5%	52	113	31.5%	5	15	25.0%	no
Epidemics	4.396	101	45	56	44.6%	29	72	28.7%	8	19	29.6%	no
Influenza Other Respir Viruses	4.380	158	72	81	47.1%	63	91	40.9%	2	4	33.3%	no
Medical Mycology	4.076	314	172	134	56·2%	115	195	37.1%	7	18	28.0%	no
J Glob Antimicrob Resist	4.035	343	174	157	52·6%	103	232	30.7%	15	44	25.4%	yes
Open Forum Infect Dis	3.835	823	382	431	47.0%	291	529	35.5%	4	9	30.8%	no
One Health	3.800	50	31	19	62.0%	20	29	40.8%	1	1	50.0%	yes
Curr Infect Dis Rep	3.725	102	57	45	55.9%	42	59	41.6%	10	9	52.6%	no
J Virus Erad	3.696	79	43	33	56.6%	28	49	36.4%	3	ő	100.0%	ves
Infection	3.553	226	83	141	37.1%	52	170	23.4%	1	15	6.3%	no
Sexually Transm Infect	3.519	210	114	92	55.3%	106	102	51.0%	18	13	58.1%	yes
Infect Dis (London)	3.404	161	67	90	42.7%	60	98	38.0%	.0	4	20.0%	no
Eur J Clin Microbiol Infect Dis	3.267	555	279	260	51.8%	173	366	32.1%	2	17	10.5%	no
Infect Control Hosp Epidemiol	3.254	449	224	223	50.1%	174	272	39.0%	4	5	44.4%	ves
Pathogens and Disease	3.166	163	80	77	51.0%	62	98	38.8%	2	14	12.5%	no
J Pediatric Infect Dis Soc	3.164	195	98	96	50.5%	88	106	45.4%	5		35.7%	no
Am J Infect Control	2.918	576	325	242	57.3%	261	307	46.0%	ğ	3	75.0%	ves
Diagn Microbiol Infect Dis	2.803	451	205	237	46.4%	148	293	33.6%	1	6	14.3%	no
Med J Hematol Infect Dis	2.576	125	59	65	47.6%	36	89	28.8%	1	7	12.5%	no
AIDS Reviews	2.500	40	9	31	22.5%	8	32	20.0%	3	3	50.0%	no
Transpl Infect Dis	2.228	337	169	163	50.9%	106	230	31.5%	2	6	25.0%	no
Surgical Infections	2.150	227	68	150	31.2%	61	164	27.1%	0	2	0.0%	no
Braz J Infect Dis	1.949	137	78	57	57.8%	48	86	35.8%	3	2	60.0%	no
Rev Inst Med Trop Sao Paulo	1.846	130	84	44	65.6%	60	70	46.2%	9	5	64.3%	
J Chemother	1.240	112	58	51	53.2%	28	83	25.2%	9	2	33.3%	yes
J Vec Borne Dis	1.688	109	45	58	43.7%	20	74	25.3%	3	2 8		yes
J vec Borne Dis Jon J Infect Dis	1.988	109	45 63	58 119	43·7% 34·6%	25 31	74 149	25.3%	3	8 21	27·3% 4·5%	no
HIV Res Clin Pract	1.362	185	63	119	34·6% 53·3%	31	149	46.7%	1	21	4·5% 33·3%	no
Leprosv Review	0.537	15	8 49	31	53·3% 61·3%	30	8 47	46·7% 39·0%	1	2	33-3%	yes
			49 57						3	5		no
J Pediatr Infect Dis - Ger	0.293	99	57	38	60·0%	40	58	40.8%	9	8	52.9%	no
TOTAL		11,027	5,350	5,503	49.3%	3,788	7,077	34.9%	190	387	32.9%	13

- Table 3. Quasi-Poisson regression model for aggregated data (ID, OB/GYN, CARDIO); EiC: editor-in-chief; IF:
- 498 499 500 501 impact factor. Statistically significant p-values are indicated in bold. R² Nagelkerke 0.998 (women first author) and 1.000 (women last author).

Variable	Incidence rate ratio	95% confidence interval	p-value				
Outcome: woman first author							
Woman EiC	1.22	1.08 – 1.39	0.002				
Woman editor proportion	1.80	1.32 – 2.44	<0.001				
Woman last author	1.26	1.09 – 1.45	0.002				
Journal IF	0.92	0.78 – 1.09	0.346				
Outcome: woman last author							
Woman EiC	1.29	1.09 – 1.52	0.003				
Woman editor proportion	2.56	1.70 – 3.85	<0.001				
Woman first author	1.28	1.05 – 1.58	0.017				
Journal IF	0.90	0.72 – 1.14	0.395				

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Table 4. Quasi-Poisson regression model for ID journals with women first author or women last author as

503 504 505 dependent variable; EiC: editor-in-chief; IF: impact factor. Statistically significant p-values are indicated in bold. R² Nagelkerke 0.700 (women first author) and 0.855 (women last author).

Variable	Incidence rate ratio	95% confidence interval	p-value				
Outcome: woman first author							
Woman EiC	1.08	0.98 – 1.18	0·123				
Woman editor proportion	1.32	1.06 – 1.63	0·012				
Woman last author	1.08	1.08 – 1.19	0·120				
Journal IF	0.91	0.81 – 1.04	0·164				
Outcome: woman last author							
Woman EiC	1.02	0.90 – 1.15	0.792				
Woman editor proportion	1.92	1.45 – 2.55	<0·001				
Woman first author	1.07	0.93 - 1.22	0.350				
Journal IF	0.97	0.82 – 1.15	0·724				

508 FIGURES

Figure 1. Correlation between woman editor proportion and woman author proportion, journal impact factor and 511 woman author proportion, and journal impact factor and woman editor proportion for each specialty category.

Figure 2. Scatter plot of journals of each specialty, each dot represents one journal. X-axis shows woman last authors per analysed articles, y-axis shows woman first authors per analysed articles.