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Data Availability Statement: Data and code to reproduce analysis are available on OSF with the following citation. Padilla BJ, Sutherland C. Defining dual-axis landscape gradients of human influence for studying ecological processes **RESEARCH ARTICLE**

Defining dual-axis landscape gradients of human influence for studying ecological processes

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Abstract

Ecological processes are strongly shaped by human landscape modification, and understanding the reciprocal relationship between ecosystems and modified landscapes is critical for informed conservation. Single axis measures of spatial heterogeneity proliferate in the contemporary gradient ecology literature, though they are unlikely to capture the complexity of ecological responses. Here, we develop a standardized approach for defining multi-dimensional gradients of human influence in heterogeneous landscapes and demonstrate this approach to analyze landscape characteristics of ten ecologically distinct US cities. Using occupancy data of a common human-adaptive songbird collected in each of the cities, we then use our dual-axis gradients to evaluate the utility of our approach. Spatial analysis of landscapes surrounding ten US cities revealed two important axes of variation that are intuitively consistent with the characteristics of multi-use landscapes, but are often confounded in single axis gradients. These were, a hard-to-soft gradient, representing transition from developed areas to non-structural soft areas; and brown-to-green, differentiating between two dominant types of soft landscapes: agriculture (brown) and natural areas (green). Analysis of American robin occurrence data demonstrated that occupancy responds to both hard-to-soft (decreasing with development intensity) and brown-to-green gradient (increasing with more natural area). Overall, our results reveal striking consistency in the dominant sources of variation across ten geographically distinct cities and suggests that our approach advances how we relate variation in ecological responses to human influence. Our case study demonstrates this: robins show a remarkably consistent response to a gradient differentiating agricultural and natural areas, but city-specific responses to the more traditional gradient of development intensity, which would be overlooked with a single gradient approach. Managing ecological communities in human dominated landscapes is extremely challenging due to a lack of standardized approaches and a general understanding of how socio-ecological systems function, and our approach offers promising solutions.

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Introduction

Rapid expansion of the global human population has led to increasing concern for natural systems and biodiversity. Anthropogenic landscape modification profoundly influences resource availability and habitat quality, which in turn, determines patterns of species distribution and abundance [1]. Given the explicit link between patterns of landscape structure and ecological processes, and the extent of human modification to the landscape, informed conservation and ecosystem management requires reliable descriptors of landscape heterogeneity gradients with an anthropogenic focus [2, 3]. Nevertheless, well documented variability in the quality, complexity, and ecological relevance of quantitative measurements of landscape structure have contributed to a lack of a general and scalable understanding of how ecological processes respond to landscape heterogeneity, particularly along gradients of human modification [3–6].

The need for ecologically meaningful measures of landscape heterogeneity (i.e., composition and configuration of landscape features) to understand drivers of ecosystem responses is well recognized [7], and over time numerous conceptual, theoretical, and applied approaches have been posited [4, 5, 8, 9]. These approaches range from the patch mosaic (fragmentation) paradigm, which, while valuable in some contexts, is arguably overly simple in heterogenous landscapes [10–12], to various metrics of patch complexity and distribution [13]. Efforts to improve the ecological relevance and realism of landscape metrics has led to the development of models thought to better represent the continuous nature of landscape heterogeneity and ecological processes by extending the patch-centered perspective to incorporate the composition of the surrounding landscape [14]. Regardless of the metrics used, successful integration of ideas in spatial ecology across systems and scales requires an improved appreciation for what landscape descriptors are measuring, and how they relate to ecological processes [15]. That is, reliable and accurate measures of landscape heterogeneity are a prerequisite to understanding patterns of ecological response across scales.

Efforts to understand and quantify ecological responses across anthropogenic gradients has resulted in some general, though equivocal, predictions about patterns of ecological response to spatial heterogeneity in human dominated landscapes. For instance, a negative relationship between species richness and human disturbance has been demonstrated in birds [16, 17], invertebrates [18, 19], plants [20, 21], and other taxa [22, 23]. Moreover, this relationship is often non-linear, with a peak in richness in areas of intermediate human modification [24, 25]. At the species level, however, responses vary, and depend on the ecology of the species in question [26–28]. While fragmentation and human population density have been linked to decreases in movement and home range size in many species [29–32], much of the literature suggests no relationship [33–35], or uncertain relationships [36, 37] between a range of ecological processes (e.g., population size, species distribution) and landscape change. These apparent contradictions suggest that measured responses to gradients of landscape heterogeneity are context or locale specific and has led to calls for improved measures of human-dominated landscapes that move towards a more general understanding of ecological dynamics in human-dominated ecosystems [6, 38].

Attempts to improve the applicability and scalability of landscape metrics used in ecological analyses has led to almost exclusive use of one-dimensional gradients of variation (e.g., percent impervious surface), even though landscape heterogeneity, and in particular the myriad ways humans alter landscapes, is multi-dimensional [39, 40]. Highly dimensional landscapes, when compressed into one-dimensional descriptors, are likely to fall short in terms of ecological realism, i.e., the landscape as perceived by a species or community, limiting the ability to infer links between landscape patterns and ecological processes, with important consequences regarding how ecosystem processes are understood and managed in the Anthropocene. We

propose an extension of the typical one-dimensional approach, which involves the identification of multiple axes of landscape heterogeneity in the context of human influence.

In this paper, we develop a multi-dimensional approach to defining landscape heterogeneity that can be used for making inferences about species distributions in human dominated landscapes. We demonstrate the generality of our multi-dimensional gradient approach by jointly analyzing urban-exurban landscapes in ten geographically and ecologically distinct US cities, identifying two significant and biologically relevant axes of variation. We demonstrate the utility of our approach in a case study analysis of American robin (*Turdus migratorious*) occupancy. Specifically, by jointly analyzing detection—non-detection data from the same ten landscapes, we investigate continental-scale consistencies in species responses to two gradients of human influence.

Methods

We selected ten geographically distinct medium sized cities (population between 200,000 and 500,000), widely distributed throughout the contiguous United States, representing the Level I ecoregions as defined by the U.S. Environmental Protection Agency [41]. These were Worcester (Massachusetts), Lexington (Kentucky), Jackson (Mississippi), Lincoln (Nebraska), Lubbock (Texas), Salt Lake City (Utah), Albuquerque (New Mexico), Bakersfield (California), Portland (Oregon), and Spokane (Washington, Fig 1). For each city, we extracted 30-m resolution landcover data from the freely available 2016 National Land Cover Database [42] for a 50-by-50 kilometer window surrounding the city center (coordinates extracted from www.latlong.net). This spatial extent extends well into exurban regions and thus represented the full extent of landscape heterogeneity for each city. To test for sensitivity to the extent, we repeated the analysis at alternative windows and found no difference in our results (*Effects spatial extent: 30 x 30 km city window* in S1 File).

Landscape composition was fairly evenly split between three dominant lands cover categories when aggregated among all cities: developed (20.03%), forests (23.98%), and agriculture (31.28%) and contained fifteen of the nineteen Anderson Land-Cover classes used by the NLCD. The remaining four ('Perennial Ice-Snow', 'Dwarf Scrub', 'Sedge-Herbaceous',



Fig 1. Map of 10 study cities. Map showing the locations of all study cities for the landscape quantification and ecological case study. Background colors represent unique Level 1 EPA Eco-Regions. Study cities are represented by numbered red points. 1—Portland, OR, 2—Bakersfield, CA, 3—Spokane, WA, 4—Salt Lake City, UT, 5—Albuquerque, NM, 6—Lubbock, TX, 7—Lincoln, NE, 8—Jackson, MS, 9—Lexington, KY, and 10—Worcester, MA. Ecoregion GIS data was sourced from the US EPA—Ecoregion spatial database (https://www.epa.gov/eco-research/ecoregions-north-america). Map was produced using the package 'map' in R.

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'Lichen') are restricted to Alaska or high elevation locations. At the individual city level, landscape composition was more variable; forested classes dominated Worcester and Spokane (39.08%, 30.22%), Albuquerque was largely scrubland (46.81%), Lexington dominated by pasture (62.31%), and agriculture in Lincoln, Bakersfield, and Lubbock (47.61%, 42.2%, 72.15%). Details for each city are provided in Table 1.

Landscape analyses followed the landscape quantification framework outlined by Padilla and Sutherland [3]. Our decisions regarding the types of landscape features relevant for analysis, the data to represent those features, and the spatial scales of analysis were made to reflect a typical ecological analysis—definitions of, and justification for, these decisions are provided in Table 2. In general, the landscapes within which our cities were set were characterized by a mosaic of natural (forests and wetlands) and un-natural (crop and developed) land-cover categories which are captured well in the NLCD classification system.

The NLCD is a 30-m resolution raster dataset where each landscape pixel is classified as a single cover type. Ecosystems are influenced both by characteristics of a fixed location, and by

Table 1. Table of study cities.

| City, State | Population | Level I Ecoregion | Open Water | Devel. | Forests | Scrub Grass | Crop Pasture |
|--------------------|------------|--------------------------------|------------|--------|---------|-------------|--------------|
| Worcester, MA | 185,877 | ER5 –Northern Forests | 3.33 | 23.15 | 65.51 | 2.54 | 6.11 |
| Spokane, WA | 208,916 | ER6 –NW Forested Mountains | 1.23 | 14.46 | 31.31 | 31.91 | 20.62 |
| Salt Lake City, UT | 200,591 | ER6 –NW Forested Mountains | 11.02 | 23.00 | 36.37 | 23.34 | 5.29 |
| Portland, OR | 583,776 | ER7 –Marine West Coast Forest | 3.12 | 37.06 | 23.43 | 7.85 | 28.54 |
| Lexington, KY | 323,780 | ER8 –Eastern Temperate Forests | 0.56 | 15.31 | 16.65 | 1.03 | 66.44 |
| Jackson, MS | 164,422 | ER8 –Eastern Temperate Forests | 4.85 | 30.11 | 43.29 | 12.16 | 18.87 |
| Lubbock, TX | 255,885 | ER9 –Great Plains | 0.15 | 12.58 | 0.23 | 14.89 | 72.15 |
| Lincoln, NE | 287,401 | ER9 –Great Plains | 1.54 | 13.02 | 5.86 | 29.4 | 50.82 |
| Albuquerque, NM | 560,218 | ER10 –North American Deserts | 0.23 | 17.66 | 15.96 | 67.71 | 3.37 |
| Bakersfield, CA | 383,679 | ER11 –Mediterranean | 0.51 | 13.96 | 1.19 | 37.93 | 46.41 |

List of ten urban-exurban regions used for landscape comparisons, including population size (2010 census) and US-EPA Ecoregion. Values for land cover types represent the percent coverage in a given city.

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Table 2. Landscape analysis decision table.

| | Decision | Justification | | | |
|--------------------------|---|--|--|--|--|
| 1) Landscape Features | Physical land-cover and demographic land-use | 'Land-cover' categories (i.e. forest, shrub) track changes in 'natural' landscapes, while 'land-use' (devel., crop) tracks the human footprint and approximate population density | | | |
| 2) Spatial Data | Remote-sensed, National Land Cover Data (2016) | NLCD land-cover data is readily available and is a consistent data-source to represent landscape features in all 10 study cities | | | |
| 3) Spatial Scale | 500-m and 1,500-m Gaussian kernel | Spatial extent (50 <i>x</i> 50-km) chosen to capture sufficient spatial and ecological heterogeneity. Primary spatial grain (500-m kernel) selected to represent breeding home range of American robin. 1,500-m as a common scale in ecological research selected to compare effects of scale. | | | |

Decisions made within landscape gradient framework for analyzing urban landscapes in jointly across study cities and in the city-specific analysis. This follows the framework outlined in Padilla & Sutherland 2019. Justification provided here is in light of dual research goals. First, to quantify landscape pattern in 10 distinct cities, and second, to evaluate occupancy patterns of American robin in response to landscape gradients.

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the local landscape context surrounding a given location [43]. Therefore, for each NLCD landcover class, we extracted a binary raster surface (1 if focal class, 0 if otherwise), and to account for landscape surrounding a given location (i.e., landscape context) we computed the spatially weighted average for each pixel using a Gaussian kernel spatial smooth, resulting in a continuous surface ranging from 0 (no focal class within smoothing kernel) to 1 (smoothing neighborhood entirely focal class). The width of the kernel defines the spatial grain of analysis, and therefore should be selected with the research specific ecological process in mind [44]. We selected a 500-m smoothing kernel for our analysis based on the typical breeding home range size of our case study focal species, the American robin [45]. We tested sensitivity of landscape quantification to this choice by replicating the analysis with a 1500-m spatial scale and found no effect of scale selection of downstream inference (*Effects of smoothing scale* in S1 File). All processing of the spatial data was conducted in R Version 3.5.3 [46] using the 'raster' [47], 'FedData'[48] and 'smoothie' [49] packages.

To identify dominant patterns of variation in these landscapes, we used Principal Components Analysis (PCA). PCA is one of several methods for summarizing a large number of potentially correlated variables into fewer uncorrelated axes of variation (others include factor analysis, non-metric multidimensional scaling, correspondence analysis), and it is particularly well suited to exploratory ordination and gradient analysis [50]. Using a matrix of classspecific smoothed landscape variables, we conducted PCA on the data for all cities combined. Dominant principal components were identified and selected based on a cumulative weight cut-off of the broken stick method, which retains components that explain more variance than would expected than dividing variance randomly among all components [51]. These were used to produce a spatially explicit gradient of habitat heterogeneity based on the resolution of the input data, where the value for each pixel in the resulting raster surface is the PCA weighted average calculated as the sum of that pixel's smoothed NLCD values multiplied by the corresponding PC weight for each NLCD value. We also conducted this analysis for each city independently in order to determine how well the combined (i.e., all cities) gradients described city-specific gradients. Output for our PCA analyses are reported in the Results section under Landscape Gradient Analysis.

Ecological case study

We evaluated the utility of multi-dimensional landscape heterogeneity gradients for ecological applications using a real-world case study. Specifically, using occupancy modelling we tested whether simultaneous consideration of multiple landscape gradients alters inferences about ecological responses relative to the traditional single-gradient approach. We analyzed American robin detection-nondetection data under an occupancy modelling framework using the gradients as covariates. We selected the American robin because it is a widespread generalist species, present in all ten focal cities, and because it is widely considered to be human-adaptive.

Robin detection histories were analyzed using a single-season hierarchical occupancy model, which estimates site occupancy probabilities while accounting for imperfect detection [52]. Stationary, complete checklists in which non-reporting of a species assumed to be non-detection, from surveys conducted from April 1st through September 30th 2018 were extracted from the *eBird* online database [53] using the R package 'auk' [54]. In this analysis, detection data from all cities were pooled in a single analysis. Because there was substantial variation in the number of eBird locations in each city (i.e., each unique fixed eBird survey site), and to improve balance and reduce regional bias in sample size, the data were randomly thinned to a maximum of 250 locations (Table 1 in S2 File).

The standard occupancy model consists of two sub-models: a logit-linear model describing site- and occasion-specific detection probability (p), which can be modelled using site- and occasion-specific covariates, and a second logit-linear model describing site-specific occupancy probability (ψ), that can be modelled using site-specific covariates. To account for variation in detection, we considered the following covariates: city (categorical factor), sampling date, and date² to allow for peaks or troughs in detection, and site-specific landscape gradient values. Sampling date was scaled (0–18) such that a one unit increase in the date variable represented 10 calendar days, which facilitated parameter interpretation and model convergence. A total of 26 possible detection models were considered, which included all additive combinations and only single interaction terms (Table 2 in S2 File). For occupancy, we included the effect of city, again as a factor, each of the site-specific landscape gradient values, and all combinations of city-gradient interactions for a total of 16 candidate models (Table 2 in S2 File).

We adopted a two-stage modeling approach whereby we fit and compared all possible combinations of detection covariates, each with the 'global' (most complex) model for the occupancy component [55]. Using Akaike's Information Criterion (AIC) to rank models, the best supported model for detection was carried over to the second stage, where we compared competing models for occupancy. Finally, the model selected for inference was validated by examining model residuals and performing goodness of fit tests. Occupancy analysis was conducted in the package 'unmarked' [56], while AIC model selection and goodness of fit tests were done using the 'AICcmodavg' package [57]. All analyses, were conducted in R Version 3.5.3 [46].

Results

Landscape gradient analysis

Principal components analysis of the combined (i.e., all cities) landscape data identified three axes of variation, explaining 37.1% of the cumulative variance in the data (Table 3). When each city was considered independently, the same three axes explained between 42.60 and 54.89% of the variance (Table 1 in S1 File), demonstrating the scalability of emergent land-scape gradients across scales. However, using the broken stick method [51], only the first two axes exceeded the 22.1% cumulative variance threshold for combined and city-specific analyses. The principal component explaining the largest proportion of data variation for the combined data (16.7%) was strongly negative for developed land-cover classes, with neutral or positive loadings for forested, open, and agricultural classes (Table 3). Developed classes are characterized by a high degree of impervious surface, buildings, and associated human population density, whereas the others are predominantly non-impervious natural (wetlands) or unnatural (pasture) landscapes. Thus, this first descriptor of landscape pattern can be interpreted as a transition from *hard* (characterized by impervious and human presence) to *soft* (unpaved natural or agricultural), which we refer to as a *hard-to-soft* gradient.

The second principal component explained 11.1% of the variation and showed a strong differentiation between the land use classes at the soft end of the hard-to-soft gradient. Specifically, this axis distinguishes between human modified but un-developed areas (cultivated croplands) and more natural areas (forests or wetlands). This axis is intuitively interpretable as a shift from modified agricultural landscapes, to un-developed natural regions, or, *brown*-to*green*. While the hard-to-soft axis does not distinguish between dominant types of *soft* landscapes, the second accounts for this variation between *brown* and *green* regions and is a valuable counterpart to component one producing a triangular distribution (Fig 2).

The third principal component explained 9.3% of the total variation and was not retained to produce a gradient surface as the cumulative variance of the first two principal components exceeded the broken stick cutoff. However, it is interesting in that like PC2, the third principal

| NLCD Layer | | Obs. Freq. | PC1 | PC2 | PC3 |
|------------|-------------------------------|------------------------|--------|--------|--------|
| | | Std.Dev. | 1.581 | 1.295 | 1.182 |
| | | Variance Explained (%) | 16.7 | 11.1 | 9.3 |
| Water | 11—Open Water | 2.70 | 0.042 | 0.030 | -0.04 |
| Developed | 21—Developed Open | 6.49 | -0.360 | 0.055 | 0.017 |
| | 22—Developed Low Intensity | 6.72 | -0.545 | 0.047 | -0.040 |
| | 23—Developed Medium Intensity | 4.78 | -0.553 | 0.015 | -0.125 |
| | 24—Developed High Intensity | 1.61 | -0.392 | -0.007 | -0.148 |
| Barren | 31—Barren Land | 0.79 | 0.039 | -0.017 | -0.116 |
| Forest | 41—Forest Deciduous | 10.17 | 0.119 | 0.469 | 0.171 |
| | 42—Forest Evergreen | 8.96 | 0.154 | 0.269 | -0.349 |
| | 43—Forest Mixed | 2.14 | 0.087 | 0.433 | 0.001 |
| Shrubland | 52—Scrub/Shrub | 10.56 | 0.151 | -0.012 | -0.554 |
| Herbaceous | 71—Grassland/Herbaceous | 10.85 | 0.140 | -0.339 | -0.282 |
| Cultivated | 81—Pasture/Hay | 11.25 | 0.053 | 0.082 | 0.446 |
| | 82—Cultivated Crops | 19.41 | 0.119 | -0.491 | 0.387 |
| Wetlands | 90—Woody Wetlands | 2.81 | 0.043 | 0.378 | 0.157 |
| | 95—Herbaceous Wetlands | 0.77 | 0.032 | 0.075 | -0.039 |

Table 3. Dominant principal component axes.

Dominant Principal Component axes. The Obs. Freq. column displays the percent composition of each land cover category in all cities combined. For each principal component, standard deviation, percent variance explained, and rotated variable loadings are displayed. Variables with a strong weight are in bold. The first two axes were selected because cumulative variance exceeded the 22.1% broken stick threshold.

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component reflected a divergence between modified and un-modified undeveloped areas. While PC2 differentiated natural deciduous and mixed forests from modified croplands, the third axis is a gradient from evergreen forests and scrub, to pastures (Table 3). Both PC2 and PC3, therefore, can be interpreted as brown-to-green in different habitat and land-use types.

Due to the ecological complementarity of the two dominant components, and our focus on highlighting the value of simultaneously considering multiple dimensions of human influence, our approach considers these axes jointly. However, it is worth noting that on their own, these gradients are analogous to traditional approaches that consider single gradients in isolation. The hard-to-soft gradient is consistent with traditional urban gradients focusing on the built environment (e.g., percent impervious surface or housing density) [3, 6, 58], or it's complement, percent forest cover. The more agricultural brown-to-green gradient, though less common in urban ecology, has been used in agro-ecological investigations [59, 60]. Our approach allows us to investigate ecological responses to both important characteristics of human influence simultaneously.

As a test of whether these axes were consistent locally and at varying spatial extents, we conducted the same analyses of NLCD data for each city independently as well as jointly using a 30x30 km window. Both city specific, and 30x30 km analyses revealed the same dominant axes of variation as the 50x50 km combined analysis. As expected, the component weights of NLCD classes and absolute values of axes differed, nevertheless, interpretation of these axes remained consistent (*Effects of smoothing scale: 1,500-m scale* in S1 File).

Ecological case study

Our robin analysis included data from a total of 1,703 sampling locations (sites) across all cities (min: 31 in Bakersfield, max: 250 in Worcester, Albuquerque, Portland, and Salt Lake City).



Fig 2. Conceptual diagram of multi-dimensional landscape gradient: A conceptual description of the triangular distribution captured by a multi-dimensional landscape definition that differentiates between urban, agricultural, and natural portions of the landscape along dual axes of variation. Hard and soft portions of the landscape are sorted along the vertical axis, while brown and green regions along the horizontal. This results in a multi-dimensional perspective where heterogeneity is maximized at the center of both axes.

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There were a total of 5,779 sampling visits across all cities, with a mean number of visits per site of 1.95 (range: 1-172, Table 1 in S2 File). The overall proportion of sites with a minimum of one observation (i.e., naïve occupancy) was 0.43, which varied by city from 0.387 in Spo-kane, to 0.482 in Bakersfield (Table 1 in S2 File).

Of the 26 detection models considered, only eight converged, largely due to the complex model structure. The AIC-top model (AIC_c wt = 1.0) included additive effects of both land-scape gradients, a quadratic effect of date, and a city by date interaction term (Table 4). In the second step, we used the best supported detection model to evaluate 16 candidate occupancy models. Here, a single model held the majority of support (AIC_cwt = 0.91, Table 4) and included the effects of both landscape gradients, city, and an interaction between city and the hard-to-soft gradient. The second model was identical to the top model apart from the inclusion of one additional term, the interaction between city and brown-to-green. Given the lack of support for the additional terms, as indicated by the model ranking [61], model evaluation and inference that follows is based on the top model. Examination of model residuals and a Chi-Square goodness of fit test showed adequate model fit.

There was a significant quadratic effect (*estimate* \pm *SE*) of survey date on detectability (-0.012 \pm 0.002), such that detection probability increased, reached a peak, and declined. Robin detection varied significantly along the *brown*-to-*green* axis, with robins more likely to be observed in more '*green*' landscapes (0.14 \pm 0.05), and showed a negative relationship with hard-to-soft, though confidence intervals included zero (-0.38 \pm 0.04). Date of peak detectability ranged from April 1st in Bakersfield (*date* = 0.0) to July 23rd in Portland (*date* = 11.3), while

| | Detection Model Structure | <u>K</u> <u>AIC</u> | | ΔAIC_c | AICwt | - <u>LogLik</u> | |
|----|---------------------------------------|---------------------|---------|----------------|--------|-----------------|--|
| 1 | $\sim city^* date + date^2 + HS + BG$ | 63 | 6544.02 | 0.0 | 1 | -3206.45 | |
| | ~ ¥ | | | | | | |
| 2 | ~ city*date+HS+BG | 62 | 6611.83 | 67.81 | 0 | -3241.44 | |
| | ~ ¥ | | | | | | |
| 3 | $\sim city^*HS$ | 60 | 6998.33 | 454.31 | 0 | -3436.84 | |
| | | 13 | 7001 75 | 457.73 | 0 | -3456.69 | |
| 7 | $\sim \psi$ | | 7001.75 | 437.73 | 0 | -5450.07 | |
| 5 | ~ date | 42 | 7038.66 | 494.64 | 0 | -3476.20 | |
| | ~ Ψ | | | | | | |
| 6 | $\sim BG$ | 42 | 7450.53 | 906.51 | 0 | -3682.13 | |
| | ~ Ψ | | | | | | |
| 7 | ~ 1 ~ ¥ | 41 | 7450.69 | 906.67 | 0 | -3683.27 | |
| 8 | ~ HS | 42 | 7452.32 | 908.30 | 0 | -3683.03 | |
| | ~Ψ Occupancy Model Structure | V | AIC | AAIC | AICast | LogLik | |
| | Occupancy Model Structure | <u><u> </u></u> | | | AICWI | - LOGLIK | |
| 1 | $\sim p \sim city^*HS+BG$ | 44 | 6530.44 | 0 | 0.91 | -3219.98 | |
| 2 | $\sim p \sim city^*(HS+BG)$ | 53 | 6535.21 | 4.77 | 0.09 | -3212.80 | |
| 3 | $\sim p \sim city^*HS$ | 43 | 6542.84 | 12.40 | 0 | -3227.23 | |
| 4 | $\sim p \sim city^*(HS^*BG)$ | 63 | 6544.02 | 13.58 | 0 | -3206.45 | |
| 5 | $\sim p \sim city + HS^*BG$ | 36 | 6567.86 | 37.42 | 0 | -3247.10 | |
| 6 | $\sim p \sim city+HS+BG$ | 35 | 6572.90 | 42.46 | 0 | -3250.67 | |
| 7 | $\sim p \sim city+HS$ | 34 | 6578.90 | 48.46 | 0 | -3254.71 | |
| 8 | $\sim p \sim city^*BG+HS$ | 44 | 6580.24 | 49.80 | 0 | -3244.88 | |
| 9 | $\sim p \sim HS^*BG$ | 27 | 6580.55 | 50.11 | 0 | -3262.81 | |
| 10 | $\sim p \sim city^*BG$ | 43 | 6581.49 | 51.05 | 0 | -3246.56 | |
| 11 | $\sim p \sim city + BG$ | 34 | 6582.03 | 51.59 | 0 | -3256.27 | |
| 12 | $\sim p \sim city$ | 33 | 6587.86 | 57.42 | 0 | -3260.23 | |
| 13 | $\sim p \sim BG$ | 25 | 6490.33 | 59.89 | 0 | -3269.76 | |
| 14 | $\sim p \sim HS+BG$ | 26 | 6591.46 | 61.02 | 0 | -3269.29 | |
| 15 | ~ <i>p</i> ~ <i>HS</i> | 25 | 6600.95 | 70.51 | 0 | -3275.07 | |
| 16 | ~ <i>p</i> ~ 1 | 24 | 6601.16 | 70.72 | 0 | -3276.21 | |

Table 4. Model selection table.

Model selection results for both detection and occupancy components of the American robin analysis based on sample size corrected AIC. K denotes the total number of parameters in the model and *AICwt* is the model weight. Detection was assessed with the global occupancy model and the best model for detection was used in all models for occupancy. Here, HS refers to the hard-to-soft gradient, while BG denotes brown-to-green.

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maximum detection probability ranged from 0.33 in Worcester, to 0.87 in Jackson (Fig 3; Table 3 in S2 File).

Robin occupancy varied by city and with both gradients. Holding both gradients at 0 (overall scaled average), robin occupancy ranged from a low of 0.47 (0.08) in Albuquerque, New Mexico to a high of 0.99 (0.074) in Jackson, Mississippi. Robin occupancy was positively related to the brown-to-green axis (0.52 ± 0.01), suggesting that robins are more likely to occur in more forested areas than in areas characterized as predominantly open or agricultural. This effect was universal across all cities. In contrast, and interestingly, direction and magnitude of the hard-to-soft gradient effect varied by city, i.e., the responses to the gradient describing the transitions from built to vegetative environments was specific to each city (Fig 4). For example, occupancy was positively associated with the hard-to-soft gradient in Spokane (1.61 \pm 0.819), but negatively associated with hard-to-soft in Worcester (-1.72 \pm 0.765).



Fig 3. Detection probability of American robin: Robin Detection probability as a function of survey date for each city predicted from the top model. Grey shaded area represents 95% confidence intervals and solid is the expected value. Date of peak detection probability varied between cities, but tended toward the start of the study period, which coincides with robin breeding behavior.

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Discussion

Analysis of spatially heterogeneous landscapes surrounding ten metropolitan regions revealed two statistically important and ecologically intuitive axes of variation, which offers an exciting alternative to the conventional one-dimensional approach to investigating ecological responses in human-dominated landscapes. Despite regional variation in landscape composition (Table 1 in S1 File) the dual-gradient approach we present here consistently distinguished between two distinct types of anthropogenic influences: a *hard*-to-*soft* gradient capturing a continuum of the built human environment, and a *brown*-to-*green* gradient capturing the human agricultural footprint (Fig 2). Our analysis shows that in addition to being fundamental properties of the landscape, considering these axes jointly provides ecological insight that would otherwise be overlooked using a single-axis approach (Table 3). This multi-dimensional perspective highlights the importance of considering the complexity of human-dominated landscapes and identifies a triangular distribution of human influence that presents an intuitive and generalizable framework for understanding patterns of ecological function and developing management strategies in human-dominated landscapes.

Landscape metrics that are adaptable to a variety of ecosystem contexts are needed to improve understanding of human-dominated ecosystems and effectively synthesize local and regional conservation efforts. Prior attempts to produce universal metrics for human footprint or urbanization have thus far failed to result in methodological consistency or broad uptake, in part due to methodological complexity and data requirements. For example the HERCULES method [62] requires users to classify the landscape into categories of building, surface cover, and vegetation using LiDAR data. Seress et al. [63] describe another method that also requires some user based classification of satellite imagery into categories of buildings, vegetation, and



Fig 4. Predicted robin occupancy along dual landscape gradients: Surface plots depicting robin occupancy on both the brown-to-green (y axis) and hard-to-soft (x axis) for each of the ten cities. The color scale ranges from low occupancy (red) to high occupancy (blue). Variation in surface plots highlights the differences in landscape composition in study cities, and the variable response to urbanization along the hard-to-soft gradient.

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road to train a semi-automated model. Metrics proposed as generalizable for use in humannatural systems also tend to focus on one axis of landscape modification, typically urbanization, rather than the full spectrum of changes to the landscape [64, 65]. Recently, a human modification gradient [66] has been produced that incorporates all aspects of the human footprint, however, it results in a single metric making it difficult to distinguish differential effects of agriculture or urbanization, as we have demonstrated here.

Multiple metrics have been used to analyze and quantify spatial change in human-dominated systems. Large suites of input variables ranging from landscape configuration measures to human population density have been used to identify multiple important features of change in urban landscapes in several notable instances [9, 65, 67, 68]. In most of these cases, however, multiple univariate measures are identified (e.g., using multi-variate analyses) as representative of landscape change along urban-rural gradients. Meanwhile, Berland and Mason [67], noted that dominant factors or principal components could perhaps be used to directly represent urbanization rather than selecting the variables with highest loading. Ultimately, regardless of the number of metrics utilized, or how they were derived, prior multi-metric research has tended to focus on identifying one aspect of landscape change, namely urbanization.

Furthermore, despite identifying multiple important measures, none of these explicitly promote a dual-axis or multi-dimensional application of these measures.

The multi-dimensional landscape gradient approach we propose here offers the flexibility to balance regional adaptability with local specificity and ecological realism to better understand more mechanistically the relationships between landscape structure and ecological processes [58, 69]. We use an established multi-variate statistical approach to succinctly describe spatial heterogeneity and employ readily available NLCD data to incorporate complexity of the entire landscape into a clear and consistent dual axis of human-influence. Although the NLCD dataset is limited to the United States, it employs a nearly identical landscape classification system as other products, including the European Space Agency's GlobCover data [70] or Copernicus Global Land Cover [71], and therefore should be applicable globally wherever such landcover data are available. In addition to PCA, other multi-variate methods have been suggested as alternatives when identifying landscape gradients, such as factor analysis [9], and non-metric multidimensional scaling [72]. More recently, PCA approaches that better account for similarities and differences in multi-group (e.g., multi-city) data have been developed [73], which may be particularly applicable when analyzing landscape gradients in multiple regions simultaneously. Our analysis has demonstrated the importance of considering multiple axes of variation in landscape gradients, and fits within a methodological framework centered on transparent and ecologically informed analysis [3]. It is important to note, however, that cityspecific means (i.e., mean effect of landscape on occupancy) in our global analysis may influence interpretation and comparison of effect size between groups. Ultimately, the multi-variate method selected by researchers should be informed by the study's goals, objectives, and types of data available [50]. As the human population continues to grow, the urban, industrial, and agricultural infrastructure must be restructured to ensure future ecological integrity, and the resulting debate over how to effectively do this has led to discussions of land-sharing, i.e., integrating natural systems into the mix of human land-uses, versus land-sparing, i.e., where natural and human systems are concentrated in large, individualized patches. Due to a traditional one-dimensional perspective of landscape heterogeneity, this discussion has largely taken place for agricultural [74], and urban [75] systems in isolation. In reality, however, urban, agricultural, and natural landscapes are inherently inter-mixed. Viewing the land-sharing versus land-sparing debate through a multi-dimensional lens of landscape heterogeneity views the landscape mosaic as a more realistic integrated agro-urban-natural system. Furthermore, the species that will benefit or suffer most from any specific sharing or sparing management, depends entirely on the landscape context within which they are evaluated [76]. Determining how to design a conservation strategy and manage a heterogeneous regional landscape for this species would require that the entire human-natural mosaic be considered and could be facilitated with a multi-dimensional approach to landscape context.

American robins are widely considered to be urban-adaptive and are thought to benefit from urbanized (e.g., *hard*) landscapes with human habitation [77, 78]. However, our results consistently predicted higher occupancy in more forested (*green*) regions over areas predominantly agricultural (*brown*), while the effect of the hard-to-soft axis on robin occupancy varied by city both in terms of magnitude and direction likely due to regional variation in composition of the *soft* landscape (Fig 4). Regional variation in the effect of hard-to-soft on robin occupancy demonstrates the need to consider and decouple multiple dimensions of landscape heterogeneity and suggests that ecological response to human-dominated landscapes is highly nuanced and regionally variable. While highly adaptable and able to exploit many habitat types, robins showed a preference for natural areas in proximity to urbanization (i.e., *green*-and-*hard*) over those in more agricultural landscapes. Our approach synthesizes prior research on the species where single landscape gradients were considered in isolation. In urban contexts higher presence and survival of robins was reported in residential yards, woodlots and golf courses [79, 80], while studies in agricultural landscapes found that robins were more common in habitat fragments surrounded by urbanization than those surrounded by agriculture [81].

City-specific variation in robin response to landscape heterogeneity reiterates the importance of landscape context on biotic response. The size (i.e., spatial extent) and density (i.e., human population) of human-dominated landscapes significantly impacts the direction and magnitude of biotic response, with larger and more densely populated cities typically resulting in a stronger negative response [82, 83]. Had our analysis centered on larger or smaller urban regions the specifics of robin response may have differed, however, our core findings-the importance of considering multiple landscape gradients and regional variation in response—would likely have remained. Though, additional research into the impacts of size of human-dominated landscapes on the use of multi-dimensional landscape gradients is warranted. We saw that robin occupancy was demonstrably influenced by both axes of human-modification across the continental United States, suggesting that a continued reliance on one-dimensional landscape descriptors may result in ecosystem pattern being misinterpreted as inherent stochasticity (e.g., noise), when in fact it reflects an overlooked component of the landscape. Specifically in our context, an analysis using a conventional hard-soft gradient would have overlooked the value of green (natural) landscapes integrated in hard (urban) regions for robins (Fig 4). Bearing this in mind, management decisions that consider only a single aspect of the human-natural landscape may overlook or misinterpret ecological response and result in ineffective conservation plans [84].

All measures of landscape heterogeneity are imperfect representations of reality and therefore fall short to varying degrees, and it is unlikely that any single metric will be ideally suited to every question of ecological pattern and process [85]. Therefore, extending one-dimensional descriptors to a multi-dimensional perspective can help move toward a more general understanding of landscape mosaics. And yet, oversimplified one-dimensional measures such as percent forest cover, or percent impervious surface continue to dominate the literature [3]. Multicity analysis of urban ecosystems has experienced rich growth in recent years. This work has highlighted the negative and positive potential impacts of urbanization on biodiversity, while stressing the importance of the regional landscape context in driving the direction and magnitude of biological response [82, 83, 86]. Still, multi-region analysis remains hampered by inconsistencies in study design and methodological limitations [87, 88]. The multidimensional, dual-axis understanding of spatial heterogeneity we describe has the potential to improve and standardize existing approaches to producing ecologically relevant landscape metrics leading to improvements in multi-region research and valuable insight into patterns of ecological response within and across human-dominated systems.

Supporting information

S1 File. Supplemental landscape gradient analysis. (DOCX)

S2 File. American robin occupancy analysis: Supplementary tables and figures. (DOCX)

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References

- Scheiner SM, Willig MR. A general theory of ecology. Theor Ecol. 2008; 1: 21–28. https://doi.org/10. 1007/s12080-007-0002-0
- 2. Turner MG. Landscape Ecology: the effect of pattern on process. Annu Rev. 1989; 232–243.
- 3. Padilla BJ, Sutherland C. A framework for transparent quantification of urban landscape gradients. Landsc Ecol. 2019; 34: 1219–1229. https://doi.org/10.1007/s10980-019-00858-z
- Wu J, Hobbs RJ. Key issues and research priorities in landscape ecology: An idiosyncratic synthesis. Landsc Ecol. 2002; 17: 355–365. https://doi.org/10.1023/A:1020561630963
- 5. Li H, Wu J. Use and misuse of landscape Índices. Landsc Ecol. 2004; 19: 389–399. https://doi.org/10. 1023/B:LAND.0000030441.15628.d6
- Moll RJ, Cepek JD, Lorch PD, Dennis PM, Tans E, Robison T, et al. What does urbanization actually mean? A framework for urban metrics in wildlife research. J Appl Ecol. 2019; 56: 1289–1300. https:// doi.org/10.1111/1365-2664.13358
- 7. Watt AS. Pattern and process in the plant community. J Ecol. 1947; 35: 1–22.
- Gustafson EJ. How has the state-of-the-art for quantification of landscape pattern advanced in the twenty-first century? Landsc Ecol. 2018;0123456789. https://doi.org/10.1007/s10980-018-0709-x
- du Toit MJ, Cilliers SS. Aspects influencing the selection of representative urbanization measures to quantify urban–rural gradients. Landsc Ecol. 2011; 26: 169–181. https://doi.org/10.1007/s10980-010-9560-4
- Frazier AE, Kedron PJ. Landscape metrics: past progress and future directions. Curr Landsc Ecol Reports. 2017; 63–72. https://doi.org/10.1007/s40823-017-0026-0
- 11. Uuemaa E, Mander Ü, Marja R. Trends in the use of landscape spatial metrics as landscape indicators: A review. Ecol Indic. 2013; 28: 100–106.
- Gustafson EJ. Quantifying landscape spatial pattern: what is the state of the art? Ecosystems. 1998; 1: 143–156. Available: https://doi.org/10.1007/s100219900011
- McGarigal K, Tagil S, Cushman SA. Surface metrics: An alternative to patch metrics for the quantification of landscape structure. Landsc Ecol. 2009; 24: 433–450. <u>https://doi.org/10.1007/s10980-009-9327-y</u>
- Eigenbrod F. Redefining landscape structure for ecosystem services. Curr Landsc Ecol Reports. 2016; 1: 80–86. https://doi.org/10.1007/s40823-016-0010-0
- Costanza JK, Riitters K, Vogt P, Wickham J, Vogt P. Describing and analyzing landscape patterns: where are we now, and where are we going? Landsc Ecol. 2019; 6. https://doi.org/10.1007/s10980-019-00889-6
- Sacco AG, Rui AM, Bergmann FB, Mueller SC, Hartz SM. Reduction in taxonomic and functional bird diversity in an urban area in Southern Brazil. Iheringia Ser Zool. 2015; 105: 276–287. <u>https://doi.org/10. 1590/1678-476620151053276287</u>
- 17. Chace JF, Walsh JJ. Urban effects on native avifauna: A review. Landsc Urban Plan. 2006; 74: 46–69. https://doi.org/10.1016/j.landurbplan.2004.08.007
- Bennett AB, Gratton C. Local and landscape scale variables impact parasitoid assemblages across an urbanization gradient. Landsc Urban Plan. 2012; 104: 26–33. https://doi.org/10.1016/j.landurbplan. 2011.09.007
- Smith JR, Schmitz OJ. Cascading ecological effects of landscape moderated arthropod diversity. Oikos. 2016; 125: 1261–1272. https://doi.org/10.1111/oik.02887
- McMullin RT, Bennett LL, Bjorgan OJ, Bourque DA, Burke CJ, Clarke MA, et al. Relationships between air pollution, population density, and lichen biodiversity in the Niagara Escarpment World Biosphere Reserve. Lichenologist. 2016; 48: 593–605. https://doi.org/10.1017/S0024282916000402

- Adhikari YP, Fischer A, Fischer HS. Micro-site conditions of epiphytic orchids in a human impact gradient in Kathmandu valley, Nepal. J Mt Sci. 2012; 9: 331–342. <u>https://doi.org/10.1007/s11629-009-</u> 2262-1
- Cavia R, Ruben CG, Virginia SO. Changes in rodent communities according to the landscape structure in an urban ecosystem. Landsc Urban Plan. 2009; 90: 11–19. https://doi.org/10.1016/j.landurbplan. 2008.10.017
- Shu X, Zhang W, Li B, Pei E, Yuan X, Wang T, et al. Major factors affecting the distribution of anuran communities in the urban, suburban and rural areas of Shanghai, China. ASIAN Herpetol Res. 2016; 7: 287–294. https://doi.org/10.16373/j.cnki.ahr.150069
- Clucas B, Marzluff JM. A cross-continental look at the patterns of avian species diversity and composition across an urbanisation gradient. Wildl Res. 2015; 42: 554–562. https://doi.org/10.1071/WR15007
- Ackley J, Carter R, Henderson R, Powell R, Muelleman P. A rapid assessment of herpetofaunal diversity in variously altered habitats on Dominica. Appl Herpetol. 2009; 6: 171–184. <u>https://doi.org/10.1163/157075408X394124</u>
- 26. Banaszak-Cibicka W, Zmihorski M. Wild bees along an urban gradient: winners and losers. J Insect Conserv. 2012; 16: 331–343. https://doi.org/10.1007/s10841-011-9419-2
- Lizée M-H, Bonardo R, Mauffrey J-F, Bertaudière-Montes V, Tatoni T, Deschamps-Cottin M. Relative importance of habitat and landscape scales on butterfly communities of urbanizing areas. C R Biol. 2011; 334: 74–84. https://doi.org/10.1016/j.crvi.2010.11.001 PMID: 21262488
- Marzluff JM. A decadal review of urban ornithology and a prospectus for the future. Ibis. 2017; 159: 1– 13. https://doi.org/10.1111/ibi.12430
- Burdett CL, Crooks KR, Theobald DM, Wilson KR, Boydston EE, Lyren LM, et al. Interfacing models of wildlife habitat and human development to predict the future distribution of puma habitat. Ecosphere. 2010; 1. https://doi.org/10.1890/ES10-00005.1
- Munshi-South J, Zolnik CP, Harris SE. Population genomics of the Anthropocene: urbanization is negatively associated with genome-wide variation in white-footed mouse populations. Evol Appl. 2016; 9: 546–564. https://doi.org/10.1111/eva.12357 PMID: 27099621
- Patrick DA, Gibbs JP. Population structure and movements of freshwater turtles across a road-density gradient. Landsc Ecol. 2010; 25: 791–801. https://doi.org/10.1007/s10980-010-9459-0
- Marsh DM, Thakur KA, Bulka KC, Clarke LB. Dispersal and colonization through open fields by a terrestrial, woodland salamander. Ecol. 2004; 85: 3396–3405. Available: http://www.jstor.org/stable/3450518
- Aida N, Sasidhran S, Kamarudin N, Aziz N, Puan CL, Azhar B. Woody trees, green space and park size improve avian biodiversity in urban landscapes of Peninsular Malaysia. Ecol Indic. 2016; 69: 176–183. https://doi.org/10.1016/j.ecolind.2016.04.025
- Cameron GN, Culley TM, Kolbe SE, Miller AI, Matter SF. Effects of urbanization on herbaceous forest vegetation: the relative impacts of soil, geography, forest composition, human access, and an invasive shrub. Urban Ecosyst. 2015; 18: 1051–1069. https://doi.org/10.1007/s11252-015-0472-6
- **35.** Dahirel M, Seguret A, Ansart A, Madec L. Dispersal-related traits of the snail *Cornu aspersum* along an urbanisation gradient: maintenance of mobility across life stages despite high costs. Urban Ecosyst. 2016; 19: 1847–1859. https://doi.org/10.1007/s11252-016-0564-y
- **36.** Hedblom M, Soderstrom B. Landscape effects on birds in urban woodlands: an analysis of 34 Swedish cities. J Biogeogr. 2010; 37: 1302–1316. https://doi.org/10.1111/j.1365-2699.2010.02299.x
- Lee M-B, Carroll JP. Avian diversity in pine forests along an urban-rural/agriculture-wildland gradient. Urban Ecosyst. 2015; 18: 685–700. https://doi.org/10.1007/s11252-014-0421-9
- McDonnell MJ, Hahs AK. The future of urban biodiversity research: Moving beyond the "low-hanging fruit". Urban Ecosyst. 2013; 16: 397–409. https://doi.org/10.1007/s11252-013-0315-2
- Cunningham RB, Lindenmayer DB. Approaches to landscape scale inference and study design. Curr Landsc Ecol Reports. 2017; 2: 42–50. https://doi.org/10.1007/s40823-016-0019-4
- Hobbs RJ. Future landscapes and the future of landscape ecology. Landsc Urban Plan. 1997; 37: 1–9. https://doi.org/10.1016/S0169-2046(96)00364-7
- 41. Omernik JM. Ecoregions: A framework for managing ecosystems. Georg Write Forum. 1995; 12: 35–50.
- 42. Yang L, Jin S, Danielson P, Homer C, Gass L, Bender SM, et al. A new generation of the United States National Land Cover Database: Requirements, research priorities, design, and implementation strategies. ISPRS J Photogramm Remote Sens. 2018; 146: 108–123. <u>https://doi.org/10.1016/j.isprsjprs.</u> 2018.09.006
- Martin AE. The spatial scale of a species' response to the landscape context depends on which biological response you measure. Curr Landsc Ecol Reports. 2018; 3: 23–33. https://doi.org/10.1007/s40823-018-0030-z

- Boyce MS, Mallory CD, Morehouse AT, Prokopenko CM, Scrafford MA, Warbington CH. Defining landscapes and scales to model landscape–organism interactions. Curr Landsc Ecol Reports. 2017; 89–95. https://doi.org/10.1007/s40823-017-0027-z
- 45. Knupp D, Owen R, Dimond J. Reproductive biology of American robins in northern Maine. Auk Ornithol Adv. 1977; 94: 80–85.
- 46. Team RC (2019) R: A language and environment for statistical computing. R Found. Stat. Comput.
- 47. Hijmans R, van Etten J. raster: Geographic data analysis and modeling. 2015.
- **48.** Bocinsky RK. FedData: Functions to automate downloading geospatial data available from several federated data sources. 2016.
- **49.** Gilleland E. Two-dimensional kernal smotthing: Using the R package Smoothie. NCAR Tech Note. 2013.
- Stuber EF, Chizinski, Christopher J, Lusk JJ, Fontaine JJ. Multivariate Models and Analyses. In: Brennan LA, Tri AN, Marcot BG, editors. Quantitative analyses in wildlife science. John Hopkins University Press; 2019. pp. 32–62.
- King JR, Jackson DA. Variable selection in large environmental data sets using principal components analysis. Environmetrics. 1999; 10: 67–77.
- Mackenzie DI, Nichols JD, Lachman GB, Droege S, Andrew J, Langtimm CA, et al. Estimating site occupancy rates when detection probabilities are less than one. Ecology. 2002; 83: 2248–2255.
- Sullivan BL, Wood CL, Bonney R, Fink D, Iliff MJ, Kelling S. eBird: A citizen-based bird observation network in the biological sciences. Biol Conserv. 2009; 142: 2282–2292.
- Strimas-Mackey MM, Hochachka WM. auk: eBird Data Extraction and Processing with AWK. R package version 0.3.0. 2018.
- Morin DJ, Yackulic CB, Diffendorfer JE, Lesmeister DB, Nielsen CK, Reid J, et al. Is your ad hoc model selection strategy affecting your multimodel inference? Ecosphere. 2020; 11. https://doi.org/10.1002/ ecs2.2997
- Fiske IJ, Chandler RB. Unmarked: An R package for fitting hierarchical models of wildlife occurrence and abundance. J Stat Softw. 2011; 43: 1–23. https://doi.org/10.18637/jss.v043.i10
- 57. Mazerolle MJ. AICcmodavg: Model selection and multimodel inference based on (Q)AIC(c). R package version 2.2–2. 2019.
- McDonnell MJ, Hahs AK. The use of gradient analysis studies in advancing our understanding of the ecology of urbanizing landscapes: current status and future directions. Landsc Ecol. 2008; 23: 1143– 1155. https://doi.org/10.1007/s10980-008-9253-4
- Kehoe L, Romero-Muñoz A, Polaina E, Estes L, Kreft H, Kuemmerle T. Biodiversity at risk under future cropland expansion and intensification. Nat Ecol Evol. 2017; 1: 1129–1135. https://doi.org/10.1038/ s41559-017-0234-3 PMID: 29046577
- 60. Yamaura Y, Royle JA, Shimada N, Asanuma S, Sato T, Taki H, et al. Biodiversity of man-made open habitats in an underused country: a class of multispecies abundance models for count data. Biodivers Conserv. 2012; 21: 1365–1380. https://doi.org/10.1007/s10531-012-0244-z
- Arnold TW. Uninformative parameters and model selection using Akaike's information criterion. J Wildl Manage. 2010; 74: 1175–1178. https://doi.org/10.2193/2009-367
- Cadenasso ML, Pickett STAA, Schwarz K. Spatial heterogeneity in urban ecosystems: Reconceptualizing land cover and a framework for classification. Front Ecol Environ. 2007; 5: 80–88. https://doi.org/10. 1890/1540-9295(2007)5[80:SHIUER]2.0.CO;2
- Seress G, Lipovits Á, Bókony V, Czúni L. Quantifying the urban gradient: A practical method for broad measurements. Landsc Urban Plan. 2014; 131: 42–50. https://doi.org/10.1016/j.landurbplan.2014.07. 010
- Luck M, Wu J. A gradient analysis of urban landscape pattern: A case study from the Phoenix metropolitan region, Arizona, USA. Landsc Ecol. 2002; 17: 327–339. https://doi.org/10.1023/A:1020512723753
- Suarez-Rubio M, Krenn R. Quantitative analysis of urbanization gradients: a comparative case study of two European cities. J Urban Ecol. 2018; 4: 1–14. https://doi.org/10.1093/jue/juy027
- 66. Kennedy CM, Oakleaf JR, Theobald DM, Baruch-Mordo S, Kiesecker J. Managing the middle: a shift in conservation priorities based on the global human modification gradient. Glob Chang Biol. 2019; 1–17. https://doi.org/10.1111/gcb.14437 PMID: 30422366
- Berland A, Manson SM. Patterns in residential urban forest structure along a synthetic urbanization gradient. Ann Assoc Am Geogr. 2013; 103: 749–763. https://doi.org/10.1080/00045608.2013.782598
- **68.** Modica G, Vizzari M, Pollino M, Fichera CR, Zoccali P, Di Fazio S. Spatio-temporal analysis of the urban–rural gradient structure: an application in a Mediterranean mountainous landscape (Serra San Bruno, Italy). Earth Syst Dyn. 2012; 3: 263–279. https://doi.org/10.5194/esd-3-263-2012

- Frazier AE, Kedron P. Landscape metrics: past progress and future directions. Curr Landsc Ecol Reports. 2017; 2: 63–72. https://doi.org/10.1007/s40823-017-0026-0
- Bicheron P, Amberg V, Bourg L, Petit D, Huc M, Miras B, et al. Geolocation assessment of MERIS Glob-Cover orthorectified products. IEEE Trans Geosci Remote Sens. 2011.
- Buchhorn M, Smets B, Bertels L, Lesiv M, Tsendbazar N-E, Masiliunas D, et al. Copernicus Global Land Service: Land Cover 100m: Collection 3: epoch 2015: Globe (Version V3.0.1). Zenodo. 2020; 1–14.
- Gu H, Singh A, Townsend PA. Detection of gradients of forest composition in an urban area using imaging spectroscopy. Remote Sens Environ. 2015; 167: 168–180. https://doi.org/10.1016/j.rse.2015.06.010
- 73. Codesido S, Hanafi M, Gagnebin Y, Gonzalez-Ruiz V, Rudaz S, Boccard J. Network principal component analysis: a versetile tool for the investigation of multigroup and multiblock datasets. Bioinformatics. 2021; 37: 1297–1303. https://doi.org/10.1093/bioinformatics/btaa954 PMID: 33165510
- 74. Phalan B, Onial M, Balmford A, Green RE. Reconciling food production and biodiversity conservation: land sharing and land sharing compared. Science. 2011; 333: 1289–1291. https://doi.org/10.1126/ science.1208742 PMID: 21885781
- Norton BA, Evans KL, Warren PH. Urban biodiversity and landscape ecology: patterns, processes and planning. Curr Landsc Ecol Reports. 2016; 1: 178–192. https://doi.org/10.1007/s40823-016-0018-5
- 76. Stott P. How climate change affects extreme weather events. Science. 2016; 352: 1517–1518. https:// doi.org/10.1126/science.aaf7271 PMID: 27339968
- Morneau F, Lépine C, Décarie R, Villard MA, DesGranges JL. Reproduction of American robin (*Turdus migratorius*) in a suburban environment. Landsc Urban Plan. 1995; 32: 55–62. <u>https://doi.org/10.1016/0169-2046(94)00177-5</u>
- Evans BS, Ryder TB, Reitsma R, Hurlbert AH, Marra PP. Characterizing avian survival along a rural-tourban land use gradient. Ecology. 2015; 96: 1631–1640. https://doi.org/10.1890/14-0171.1
- Malpass JS, Rodewald AD, Matthews SN, Kearns LJ. Nest predators, but not nest survival, differ between adjacent urban habitats. Urban Ecosyst. 2018; 21: 551–564. https://doi.org/10.1007/s11252-017-0725-7
- **80.** Blair RB. The effects of urban sprawl on birds at multiple levels of biological organization. Ecol Soc. 2004; 9.
- Rodewald AD, Bakermans MH. What is the appropriate paradigm for riparian forest conservation? Biol Conserv. 2006; 128: 193–200. https://doi.org/10.1016/j.biocon.2005.09.041
- Beninde J, Veith M, Hochkirch A. Biodiversity in cities needs space: a meta-analysis of factors determining intra-urban biodiversity variation. Ecol Lett. 2015; 18: 581–592. https://doi.org/10.1111/ele.12427 PMID: 25865805
- Fidino M, Gallo T, Lehrer EW, Murray MH, Kay CAM, Sander HA, et al. Landscape-scale differences among cities alter common species' responses to urbanization. Ecol Appl. 2021; 31: 1–12. https://doi. org/10.1002/eap.2253 PMID: 33141996
- Fischer JD, Lindenmayer DB, Kajtala V. Beyond Fragmentation: The continuum model for fauna research and conservation in human-modified landscapes. Oikos. 2006; 112: 473–480.
- Frazier AE. Emerging trajectories for spatial pattern analysis in landscape ecology. Landsc Ecol. 2019; 1. https://doi.org/10.1007/s10980-019-00880-1
- Aronson MFJ, La Sorte FA, Nilon CH, Katti M, Goddard MA, Lepczyk CA, et al. A global analysis of the impacts of urbanization on bird and plant diversity reveals key anthropogenic drivers. Proc R Soc B Biol Sci. 2014; 281: 20133330–20133330. https://doi.org/10.1098/rspb.2013.3330 PMID: 24523278
- Magle SB, Fidino M, Lehrer EW, Gallo T, Mulligan MP, Ríos MJ, et al. Advancing urban wildlife research through a multi-city collaboration. Front Ecol Environ. 2019; 17: 232–239. <u>https://doi.org/10.1002/fee.</u> 2030
- Knapp S, Aronson MFJ, Carpenter E, Herrera-Montes A, Jung K, Kotze DJ, et al. A research agenda for urban biodiversity in the global extinction crisis. Bioscience. 2021; 71: 268–279. <u>https://doi.org/10. 1093/biosci/biaa141</u>