# 1 Line Transect Sampling of Primates: Can Animal-to-Observer

## **2 Distance Methods Work?**

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14 **Short title:** Use of AODs in line transect sampling

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## 19 Line Transect Sampling of Primates: Can Animal-to-Observer

## **Distance Methods Work?**

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Abstract Line transect sampling is widely used for estimating abundance of primate populations. Animal-to-observer distances (AODs) are commonly used in analysis, in preference to perpendicular distances from the line. This is in marked contrast with standard practice for other applications of line transect sampling. We formalize the mathematical shortcomings of approaches based on AODs, and show that they are likely to give strongly biased estimates of density. We review papers that claim good performance for the method, and explore this performance through simulations. These confirm strong bias in estimates of density using AODs. We conclude that AOD methods are conceptually flawed, and that they cannot in general provide valid estimates

**Keywords** animal-to-observer distances • distance sampling • estimating primate density

• Kelker strip • modified Kelker method • primate surveys

### Introduction

of density.

- Line transect sampling is a 'distance sampling' method (Buckland et al., 2001, 2004),
- 42 widely used for estimating the abundance of wild animal populations. In most areas of
- 43 application, disciplines have standardized their methods and use the software package

Distance (Thomas *et al.*, in press). However, methods that in other disciplines are generally considered to be obsolete are still often used and recommended in primatology: the Kelker strip (Kelker, 1945) and the 'modified Kelker method' (Struhsaker, 1981), which covers a range of methods based on assessing the effective width of the searched strip from animal-to-observer distances or AODs. In addition, survey design issues are often ignored, and the precision of abundance estimates is often not quantified, compromising studies designed to compare the performance of different methods.

In this paper, we first consider strip transect sampling, and the assumptions under which it is effective. We then explore AOD methods that are conceptually related to strip transect sampling. Plumptre and Cox (2006) noted that such methods have no mathematical basis; here we show that they are based on an erroneous interpretation of the AOD distribution. We review studies that claim good performance of the approach, and assess its performance using simulation.

## **Strip Transect and Related Methods**

59 Standard Strip Transect Sampling

In standard strip transect sampling (Buckland *et al.*, 2001), we place lines at random in the survey region, or more commonly, we randomly superimpose a set of equally-spaced lines on the survey region. An observer walks along each line, recording all animals within a distance *w* of the line, where *w* is the strip half-width. Given random placement of an adequate number of lines through the survey region, this density estimate is representative of the whole survey region, allowing abundance within that region to be estimated.

Many animals, including primates, tend to occur in groups, termed 'clusters' in the distance sampling literature. In strip transect sampling, we have two main options for dealing with groups. The first is to ignore them; all animals within a sampled strip are

- 70 counted, without regard to the groups. Thus for groups that extend beyond the survey
- strip, some animals are counted and some not. The second option is to count the whole
- group if its centre is within the sampled strip, and not if its centre is beyond the strip.

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- 74 Assumptions
- 75 If groups are ignored, the key assumptions are:
- 1. Animals that are located within a sampled strip prior to any response to the observers
- are certain to be detected and counted.
- 78 2. Animals that are located outside the sampled strips are not counted.

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- 80 If groups are the recording unit, the key assumptions are:
- 1. Groups whose centres are within a sampled strip prior to any response to the observers
- are certain to be detected and counted.
- 2. The size of each of these groups is recorded without error.
- 3. Groups whose centres are outside of the sampled strip are not recorded.

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- In either case, we also assume that there is an adequate sample of randomly-
- 87 distributed strips, or a grid of strips randomly positioned, in the survey region. This
- 88 assumption is not usually listed, because it is an aspect of survey design, and survey
- 89 design is under our control: if we use an appropriate design, we guarantee that the
- 90 assumption is met. However, non-randomized designs (e.g. transects along pre-existing
- 91 trails) with inadequate replication (fewer than 10 lines; Buckland et al. (2001:232)
- 92 recommend at least 10-20 lines) are frequent in primate surveys, so we state the
- 93 assumption explicitly here. In practice, we usually prefer systematic random designs to
- 94 designs in which each transect is independently randomized.

#### **Problems**

If individual animals are counted, we can seldom be sure of detecting all animals within the sampled strips. Even if this is possible, it can be very difficult to determine whether a detected animal is within the strip, especially for animals close to the edge of the strip. If groups are recorded, it is generally difficult to estimate the location of the group centre. For these reasons, it has become standard practice amongst some survey teams to record the position of the group as being at the location of the first-detected animal (Struhsaker, 1981; Hassel-Finnegan *et al.*, 2008). This animal is more likely to be within the sampled strip than a randomly selected animal, and consequently, the strategy leads to positive bias in density estimates. This bias is substantial if average group spread is of similar magnitude to the strip half-width *w*.

Line transect sampling (Buckland *et al.*, 2001) relaxes the assumption that all groups in the strip are detected, but generates similar bias in density estimates if the location of the group is taken as the location of the first-detected animal. This source of bias is well-known (e.g. Whitesides *et al.*, 1988), yet the practice persists, and as a consequence, standard line transect sampling is often considered to overestimate density in the primate literature (Hassel-Finnegan *et al.*, 2008). Buckland *et al.* (in review) discuss how to implement standard line transect methods for primates.

#### The Kelker Strip

The Kelker strip (Kelker, 1945) is a variation on strip transect sampling. Shortest distances of detected animals from the line (so-called perpendicular distances) are recorded, as for line transect sampling. These distances are placed into intervals, and plotted in a histogram, from which the distance out to which all animals are detected is assessed. When used for primate surveys, groups are recorded, together with the distance of each group from the line.

123 Assumptions

The assumptions of this approach are essentially the same as for strip transect sampling, although we now estimate the strip half-width from the distribution of distances from the line, which requires accurate estimation of distances to the centres of detected groups, including those groups that are detected beyond the strip.

#### Problems

The method shares with strip and line transect sampling the difficulty of identifying the location of group centres. For strip transect sampling, this problem can be avoided if it is possible to record all individuals in the strip, and accurately determine that they are in the strip. However, because the Kelker strip requires distances from the line to be recorded, the distance of each detected animal from the line must be recorded to implement this approach. When groups are recorded, and distances from the line are taken as the distance of the first animal detected from the line, the method is prone to exactly the same upward bias as strip and line transect sampling.

The method also has problems that line transect sampling does not. First, it is subjective. Identification of the distance up to which all groups are detected can vary between analysts and between different choices of interval cutpoints for the histogram. Second, variance estimates ignore the uncertainty in estimating this distance, and so tend to be underestimates. Third, many observations are discarded because many animals are detected beyond the distance at which detection can be assumed certain, reducing precision. Fourth, for small sample sizes, the method tends to give biased estimates of abundance. When sample size is small, sampling variation tends to be large, and the choice of cutpoint can be influenced by chance variation in the proportion of detections close to the line. If by chance several groups are detected close to the line, there is a

tendency to set the cutpoint too small, which can lead to positive bias (Fig. 1). However, if groups are missed whose centres are inside the selected cutpoint, negative bias will occur. It is difficult to ensure a balance between these biases.

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#### AOD Methods

AODs (called 'radial distances' in the line transect literature) are often easier to measure than perpendicular distances from the line. In this case, observers generally also record the sighting angle, allowing the perpendicular distance to be calculated (Buckland et al., 2001:5). However, line transect methods that model the AOD rather than the perpendicular distance have a long history (e.g. Hayne, 1949). They are now seldom used, as the models are not plausible representations of the detection process (Hayes and Buckland, 1983). Despite this failing, they were mathematically coherent models. Unfortunately, this is not true of the AOD models used by primatologists. These have their origins in surveys conducted in Kibale Forest, Uganda by Struhsaker (1975). Initially during surveys carried out between 1970-1972, he plotted positions of monkey groups on maps and from these he calculated the perpendicular distance of each group from the line. It is not stated what he plotted on the maps: nearest animal, first animal detected, or group centre. In later surveys in 1974-76, he estimated the AOD by eye. The density estimates obtained using the perpendicular distances tended to overestimate the 'known' primate density based on knowledge of group sizes and home range of habituated groups (Struhsaker, 1975). This appears to be because perpendicular distance from the line to the nearest animal of the group was recorded, as nearly 40% of groups were recorded at zero distance (suggesting that nearly 40% of detected groups straddled the line). He then arbitrarily chose to calculate the mean and maximum AOD across detected groups, where AOD for a group was the distance to the first detected animal, and used these as estimates of the strip half-width for a Kelker strip analysis. The mean AOD tended to overestimate density of primates while the maximum AOD tended to underestimate density. He then defined a 'maximum reliable AOD' – the distance at which the frequency of sightings falls when plotting AOD against number of sightings. This too was used as an estimate of the strip half-width. In each case, only AODs less than the estimated half-width were included in the density estimate (Struhsaker, 1981). We can find no published results that show that he compared methods based on perpendicular distances and AODs measured in the field.

#### Assumptions

Beyond the assumption that the selected truncation distance results in a complete count of primate groups within that same distance of the line, assumptions are never explicitly stated for the modified Kelker method and its AOD variants. In fact, there is no coherent framework under which the methods can be justified, so that it is not possible to specify a full set of assumptions, as will be seen below.

#### Problems

The method, like the Kelker strip, is subjective when the 'maximum reliable AOD' is used, so that different analysts may select different truncation distances, and estimation is sensitive to the choice. When there is a subjective element in the analysis, and estimation is sensitive to the subjective choice, it is good practice for assessments of the performance of the method based on populations with known density to be performed blind – that is, the analyst should be unaware of the true density when generating estimates.

Authors who use AOD methods appear not to estimate the precision of their estimates. Of the papers reviewed below that use the method, only Fashing and Cords (2000) quoted standard errors, and those were based on repeat surveys of a single line, so

do not reflect spatial variation (so-called pseudo-replication, Hurlbert, 1984). Variance could be estimated as for strip transect sampling, although this fails to incorporate uncertainty in estimating the strip width.

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However, there is a more serious problem with these methods, in that their conceptual framework is erroneous. The methods confuse a probability density function When a histogram is plotted, showing frequencies of with a detection function. detections by distance intervals, then the histogram, if rescaled so that the total area of the histogram bars is unity, provides an empirical estimate of a probability density function: it shows the relative frequencies of detections by distance. By contrast, the detection function is the probability of detecting a group, as a function of distance of that group from the line or, for AOD methods, from the observer. When perpendicular distances from the line are used, the two functions have the same shape (Buckland et al., 2001:53), so that the histogram may be used for example to assess the perpendicular distance at which probability of detection starts to fall. However, if AODs are used, this is no longer the case (see e.g. Buckland et al., 2001:148). The point at which frequencies start to fall does not correspond with the point up to which probability of detection is certain. To illustrate this, we simulated data using the hazard-rate model  $g(y) = 1 - \exp\{-(y/20)^{-2}\}$ for the detection function (Fig. 2), where g(y) is the probability of detecting an animal group whose centre is at perpendicular distance y from the line. Groups had mean size of three, and half the group spread was 10m. (Full details are given in the simulation study section below.) We show a histogram of simulated AODs (Fig. 3). The 'maximum reliable AOD' might be taken as 30m or 40m, depending on the judgement of the analyst, but detectability starts to fall away at around 10m (Fig. 2). The effect is substantial; over 60% of groups are undetected at 30m, and nearly 80% at 40m. In other words, for these

values of the maximum reliable distance, we can expect underestimation of density by over 60% (30m) or nearly 80% (40m).

This flaw in the method is self-evident if you consider what is being done. For example, Hassel-Finnegan (2008) report in their Fig. 2 an estimate of 55m, estimated from AODs, up to which detection is considered to be certain for white-handed gibbons *Hylobates lar carpenteri*. However, of 155 detections, 141 of these were not detected until they were closer to the observer than 55m. Indeed, the median AOD is under 35m. If animals at 55m were certain to be detected, then AODs of less than 55m should not be observed – as the observer approaches an animal, then it will be detected at the certain detection distance of 55m, if not at a greater distance.

Marshall *et al.* (2008), while acknowledging that the method lacks a mathematical basis, state erroneously: 'Because sighting distance is used rather than distance to transect, the pattern of decline with distance is a true detection function.' This is not the case (Fig. 3). Hence their belief that the method should be used when other methods fail lacks credibility.

There is a further inconsistency in the method, when the histogram is used to identify the distance up to which detection is certain. AODs are used to estimate (erroneously) this distance. However, this is then assumed to be the half-width of a strip centred on the line, rather than the radius of a circle centred on the observer. It is also used to truncate detections whose AODs are larger. Suppose we use 55m as the truncation AOD as in Hassel-Finnegan *et al.* (2008). A group that is detected when still 80m away, but which is located on the line, is therefore excluded from the count – but its location is right at the centre of the strip to which the count supposedly relates.

If data quality were otherwise good, each variation of the method would clearly undercount groups within the sampled strip. If by chance the method does produce a good estimate, it may be a consequence of defining group location with respect to the

observer to be the location of the first detected animal of a group: the upward bias generated by this strategy might cancel with the downward bias of the modified Kelker method. There is no assurance that the biases will cancel in general.

## **Review of Papers that have Assessed AOD Methods**

Struhsaker (1981) proposed use of the modified Kelker method on the basis that it gave rise to the least biased estimates of density of red colobus monkeys *Piliocolobus oustaleti*. However, the reason for overestimation of density in his study is evident from the following quote: '... nearly 40% of the 166 sightings of red colobus were over the census transect and were scored as zero meters from the trail ...' He does not clarify how distances were measured in the field, but as all groups 'over the census transect' were recorded as being at zero distance, we can infer that distance of the nearest animal to the transect was recorded, with predictable overestimation of density; any attempt to salvage density estimates from such poor distance data will inevitably be subjective and *ad hoc*. (Another possibility is that the position of the line was not well-defined, so that any animal or animal group that was close to the line was simply recorded as on the line.)

Defler and Pintor (1985) assessed the performance of the three modified Kelker methods based on mean AOD, maximum reliable AOD and maximum AOD, against known densities of three species (red howler monkey *Alouatta seniculus*, collared titi monkey *Callicebus torquatus* and brown capuchin *Cebus apella*) in Colombia. Their design comprised a single, non-random transect; lack of randomization means that we cannot be confident that the density along the line is representative, and lack of replication means that there is no basis for assessing precision of estimates. Their results were very mixed. Mean AOD led to estimated biases of +12%, -30% and +538% for the three species. Maximum reliable AOD gave estimated biases of +12%, -12% and +668%. Maximum AOD resulted in estimated biases of -25%, -56% and +226%.

Chapman et al. (1988) used just 5 transects, subjectively placed. Their known populations comprised just a single group of each of two species (white-headed capuchin Cebus capucinus and mantled howler monkey Alouatta palliata) in Costa Rica. They used six different methods of measuring distances: 'the mean, maximum and reliable perpendicular distance from the transect to the animal first sighted and the mean, maximum and reliable distance from the observer to the animal.' Thus all six methods were prone to bias by assuming that the first animal sighted was at the centre of the group. The authors did not quantify the precision of their estimates, and did not define what a 'reliable' distance is (there is not a unique definition of it in the literature). Their estimates show poor performance of all methods, with no clear winner, yet they come down heavily in favour of methods based on AODs on the grounds that 'sightings that occur directly over the transect or at a steep angle to it, are likely to cause bias.' They do not clarify why. They also claim that the ability of the observer to estimate perpendicular distance will be limited when the terrain is rough, which in our view is not a compelling reason for using the wrong distance. Analyses presented in the paper do not in fact support the use of AOD methods; rather, misunderstanding of the methods has resulted in their recommendation to use it.

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Brugiere and Fleury (2000) did not attempt standard line transect analysis because they had only 23 detections, yet they considered this sample size to be adequate for ten other methods of analysis. Their known population comprised just three groups of a single species (black colobus *Colobus satanas*) in Gabon. Their design comprised just two transects, subjectively positioned, and they did not quantify precision of estimates. They used strip counts, with strip half-width pre-set at three values: 60, 80 and 100m; strip counts, with half-width estimated from the data, using maximum, mean or 'maximum reliable' perpendicular distance; the modified Kelker method, using maximum, mean or 'maximum reliable' AOD; and estimation of the effective strip half-

width from a histogram, coupled with adding half the group spread to this distance (Whitesides *et al.*, 1988). Uncertainty over true density complicated assessment of the methods, and they drew no firm conclusions on which method was best.

Fashing and Cords (2000) analysed data on two species (black-and-white colobus *Colobus guereza* and blue monkey *Cercopithecus mitis*) in Kenya. They estimated true densities based on home range data primarily on five groups and three groups respectively, although data from additional groups were also used. The 'design' was of a single non-random transect, placed along trails. They estimated precision from variation in repeat runs along the same transect. They estimated transect width using a) the maximum reliable AOD; b) the maximum reliable perpendicular distance; and c) the maximum reliable perpendicular distance with the addition of half the group spread, as recommended by Whitesides *et al.* (1988). They also used the shape-restricted estimator of Johnson and Routledge (1985) (a type of perpendicular distance detection function estimator that is seldom used). For both species, the method based on perpendicular distances, together with the half-group spread correction, gave estimates closest to the true density. The shape-restricted estimator performed particularly poorly.

Hassel-Finnegan *et al.* (2008) used just a single transect to estimate densities of two species (white-handed gibbons and Phayre's leaf monkeys *Trachypithecus phayrei crepusculus*) in Thailand. They assessed true density largely on the basis of a single group for each species, and did not quantify precision of estimates. They used both the Kelker strip and the modified Kelker method, with truncation distance estimated as the point at which frequencies in the respective histograms of perpendicular distances and of AODs started to fall. In addition, they used Distance (Thomas *et al.*, in press) to perform a standard line transect analysis of perpendicular distances. All detection distances, whether perpendicular distances or AODs, were measured to the first animal detected of the group. For both species, Distance and the Kelker strip gave rise to overestimates of

the true density, while the modified Kelker method gave estimates very slightly under the true density. However, given the lack of replication (a single line, and a single group of each species), it seems that little can be inferred from these results. Hassel-Finnegan *et al.* (2008) quote the papers of Chapman *et al.* (1988) and Fashing and Cords (2000) to support their contention that analyses based on AODs closely match true densities, while those based on perpendicular distances overestimate. However, the results of neither paper support this conclusion.

All of the above comparisons are based on studies where true density is established by studying a small number of habituated groups, and estimating the size of their home range. There are several reasons why there might be bias in these 'true' densities. For example home ranges of groups may partially overlap, and because the transects in these studies are positioned subjectively, they may sample parts of the home range that are favoured or avoided by the habituated group, leading to a mismatch in the densities being estimated by the two approaches. This is exacerbated when the sampled strip(s) extend beyond the home range(s) of the habituated group(s), into other home ranges. Further, lone males are not included in densities obtained from home range studies, so that density might be expected to be lower as assessed by this method than that obtained by appropriate application of line transect sampling methods. In the case of a population of grey-cheeked mangabeys *Lophocebus albigena* in Uganda (Olupot and Waser, 2005), Olupot (pers. comm.) estimates that around 30% of males are solitary, corresponding to around 8% of the total population.

Mitani et al. (2000) compared censuses of seven species (black-and-white colobus, Pennant's red colobus *Procolobus pennantii*, baboons *Papio anubis*, blue monkeys, grey-cheeked mangabeys, red-tailed monkeys *Cercopithecus ascanius* and chimpanzees *Pan troglodytes*) made along almost the same census route between years at Ngogo in Kibale Forest. The three authors measured AOD separately for their respective

census periods (1975-76 and 1996 Struhsaker; 1997-98 Lwanga; and 1996 Mitani). They also used a single transect, which formed the shape of a square route. The authors found great variation in the estimation of AOD between the three of them, showing that each observer would estimate a very different maximum reliable sighting distance and that the shape of the sighting distributions differed significantly. They therefore could not use the modified Kelker method to compare densities between years and resorted to comparing encounter rates of primate groups per kilometre walked. Variation between observers with the modified Kelker method eliminates any possibility of comparison unlike standard perpendicular distance methods where a probability of detection can be computed for each observer to allow comparisons to be made (Marques *et al.*, 2007).

Marshall *et al.* (2008) review four methods: strip transects, the modified Kelker method, and two methods based on perpendicular distances – distance from the line of the group centre, and distance from the line of the centre of measurable individuals. They note that methods based on truncating a substantial proportion of data need larger sample sizes for comparable precision than methods that do not. They also note that the assumptions of the modified Kelker method are 'unknown'. However, they still advocate its use for when expertise is unavailable to apply other methods, or when the data (location of group centres, or reliable estimates of mean group spread) or assumptions (certain detection on the line, accurate measurement to individuals) required by other methods are unachievable. They justify this recommendation on the grounds that the method consistently performs well in field trials, a claim that does not stand up to close scrutiny, as noted above. The recommendation presupposes that the *ad hoc* method will produce useful estimates when the other methods do not. The observer effects found by Mitani *et al.* (2000) for example suggest that this is unlikely.

## **Simulation Study**

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380 Simulating Populations and Samples

To assess how different methods perform, we simulated data from populations of known density. This is intentionally an idealized study, with a large sample of lines systematically spaced with a random start, with certain detection of animals on the line, no responsive movement, and no measurement error in distances. If methods perform poorly here, they can certainly be expected to in real studies. For simplicity, we assumed a rectangular survey region, 20km long and 5km wide. We placed 25 transects in the region, each 5km long, spaced 800m apart.

True number of groups in the survey region was 500, randomly spread through the region with a uniform density. Mean group size was 3, 10 or 30 animals, so that total population size was 1500, 5000 or 15000, corresponding to 15, 50 or 150 animals per square kilometre. We assigned the animals to the 500 groups by first assigning a single animal to each group. We then generated a random number for each remaining animal from a continuous uniform distribution on  $(0.500^p)$ , with p = 0.75. We raised this number to the power 1/p, and rounded up to the next integer; the resulting value defined the group to which the animal was assigned. This ensures greater variation in group size than would occur if all groups had the same expected size (corresponding to p=1), but the expectation of mean group size was 3, 10 or 30, as required. We assigned the position of each animal in a group at random within a circle of radius  $\rho$ , centred at the assigned group location, with  $\rho = 10$ , 25 and 50m. All group centres fell within the survey region, but individual animals could be assigned a location outside the survey region. To avoid the complication of partial sampling of groups straddling the boundary, we extended sampling into a bufferzone, to allow the whole group to be sampled. We did not count effort (i.e. length of transect) in the bufferzone; this does not create bias because the additional sightings compensate for the 'missing' sightings that would have occurred had groups been simulated whose centres were outside the study region, but which straddled the boundary.

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Hayes and Buckland (1983) developed a hazard-rate model of the detection process. Their model is useful here to simulate the detection process as the observer approaches a group of animals. In this study, we initially simulated whether or not an animal was detected independently of other animals in a group. We assumed a hazard function of the form  $k(r) = ar^{-b}$ , with b = 3 and b = 5, where r is distance between the animal and the observer. If the observer has not yet detected an animal at distance r, then k(r)dx is the probability that the animal is detected as the observer advances a small distance dx along the line. Given the above form for k(r), we can derive the detection function g(y), which is the probability that an animal at distance y from the line is detected:  $g(y) = 1 - \exp\{-(y/c)^{-(b-1)}\}$ . We chose (c = 20, b = 3), for which a = 400, and (c = 30, b = 5), for which a = 1215000. These two detection functions are shown (Fig. 2). To mimic the enhanced probability of detecting animals in a group once the first animal of the group has been detected, we identified all groups for which at least one animal was detected, and simulated a second 'pass' to search for undetected animals in the group, again using a hazard-rate detection function, but with the scale parameter c increased by 50% (c = 30 for scenarios with b = 3, and c = 45 when b = 5).

For a detected animal, we recorded both the AOD at the time of first detection and the perpendicular distance from the line. We did not record animals further than 150m from the line. Sample sizes were typically in the range 60-120.

- 427 Estimating Densities
- 428 For each combination of mean group size, group spread, true density and detection
- 429 function, we simulated 100 populations, and surveyed each once. We applied the
- 430 following analysis methods.

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- 1. The modified Kelker method, based on mean AOD, where AOD for each detected
- group is the distance of the first detected animal from the group. We took the mean AOD
- as an estimate of the strip half-width, and the mean of recorded sizes of detected groups
- as an estimate of mean group size in the population.
- 436 2. The modified Kelker method, based on maximum AOD, where AOD for each
- detected group is the distance of the first detected animal from the group. We took the
- maximum AOD as an estimate of the strip half-width, and the mean of recorded sizes of
- detected groups as an estimate of mean group size in the population.
- 3. The modified Kelker method, based on maximum reliable AOD, where AOD for each
- detected group is the distance of the first detected animal from the group. We grouped
- 442 AODs into 10m bins, and estimated the half-width of the strip by starting at the bin
- closest to the line (0-10m), and identifying the first bin for which the count was at most
- one half of the mean count for preceding bins. If for example the mean count in the first
- four bins was 10.5, and the count for bin 5 (40-50m) was 5, then the strip half-width was
- taken to be 40m, and detections at a greater distance were excluded from the analysis.
- We estimated mean group size in the population by the mean of recorded sizes of
- detected groups.

- 450 Results
- 451 All three modified Kelker methods have strong negative bias for all scenarios (Table 1).
- 452 The bias is consistent across different group sizes and spreads, but differs markedly by

detection function. This finding is consistent with the finding by Mitani *et al.* (2000), that density estimates were not comparable across observers. The bias is especially large for method 1, the maximum AOD method (-90.7% and -75.7% for the two detection functions). For method 2, bias was -42.3% for the first detection function and -25.1% for the second. The corresponding values for method 3 were -43.9% and -28.2%.

These biases are not fully explained by bias in recorded group sizes (Table 2). Interestingly, although bias in recorded group size increases both with mean group size and with group spread, for methods 2 and 3, bias in density estimates within a method and detection function is largely independent of mean group size and group spread. However, as the bias is not consistent across different detection functions, it suggests that neither method gives a reliable estimate of relative density.

The bias is also not attributable to recording distances to the first detected animal, rather than to the group centre. Using measurements to group centres, AODs would increase, resulting in larger estimated strip widths, and reduced densities, so that bias would be even larger.

Because the methods have no coherent mathematical framework, it is not possible to identify the causes of bias, as there are no coherent assumptions that we can assess.

### **Discussion**

In our simulation study, we found serious biases with estimators based upon AODs. Hassel-Finnegan *et al.* (2008) criticize conventional line transect sampling as implemented in Distance because a large number of detections is needed for reliable analysis. They fail to note that this is even more true of the Kelker strip, for which many of the observations are discarded (Marshall *et al.*, 2008). With inadequate sample sizes, choice of truncation distance is more subjective, uncertain and influential.

For methods based on selecting a single animal, and using the distance to it as the distance to the group centre, there is some ambiguity in the literature about whether the selected animal is the first animal detected or the closest animal. In general, the two are not the same individual. In our simulations, we assumed that it is the first animal detected. Struhsaker (1981) recorded 40% of detected groups as being on the line, which suggests that he used the distance of the closest animal to the line. Alternatively, if his transect was along trails, it may be that animals directly above the line were the first to be detected, because they were more visible.

We conclude that AOD methods as used by primatologists are conceptually flawed; the resulting estimates should not be treated as estimates of absolute density. Whether they are acceptable estimates of relative density depends on many factors. Estimates are unlikely to be comparable across different observers (Mitani *et al.*, 2000) or habitats for example. Estimating primate abundance is often difficult compared with many other taxa, as the animals often reside in hard-to-access, low-visibility areas and are often clustered, cryptic and highly mobile. Nevertheless, more reliable estimates of abundance are potentially possible by combining good survey design with better field and analytic methods (Buckland *et al.*, in review).

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### Figure legends

- Fig. 1. Shown here are two datasets, both generated from a detection function with certain detection out to 40m, and rapidly declining detection probability at larger distances. When by chance there are more detections close to the line, visual inspection of the data leads to selection of a smaller cutpoint for the Kelker method; 20m in this example. When there are more detections close to 40m, the cutpoint is likely to be set at 40m. If we fix the cutpoint in advance, at either 20m for both analyses or 40m, we expect unbiased estimates of density, but if we use 20m for the first analysis and 40m for the second analysis, we overestimate density on average. Dashed lines: mean count with truncation at 40m. Dotted lines: mean count with truncation at 20m.
- Fig. 2. The detection functions used in the simulation study. Note that these detection functions apply to each individual animal; the probability that at least one animal of a group will be detected will be larger than shown here substantially so for large groups. The solid line is  $g(y) = 1 \exp\{-(y/20)^{-2}\}$  and the dashed line  $g(y) = 1 \exp\{-(y/30)^{-4}\}$ .
- Fig. 3. Histogram of AODs simulated from the hazard-rate model of the detection process,  $g(y) = 1 \exp\{-(y/20)^{-2}\}$ .

Table 1. Mean (standard deviation in parentheses) of density estimates for the three methods of estimation.

Mean group size	3			10			30			
Half-group spread	10m	25m	50m	10m	25m	50m	10m	25m	50m	
True density	15	15	15	50	50	50	150	150	150	
$g(y) = 1 - \exp\{-(y/20)^{-2}\}:$										
Maximum AOD	1.6	1.6	1.6	4.2	4.2	4.3	12.8	13.2	13.1	
	(0.7)	(0.7)	(0.6)	(0.6)	(0.6)	(0.7)	(1.3)	(1.3)	(1.1)	
Mean AOD	8.6	8.5	8.3	29.6	29.2	29.0	85.5	88.2	87.4	
	(1.9)	(1.7)	(1.5)	(5.2)	(4.0)	(4.0)(	13.4)	(12.7)	(11.9)	
Max reliable AOD	8.8	8.4	8.4	28.7	28.2	27.6	80.9	83.4	83.7	
	(2.1)	(1.7)	(1.6)	(5.3)	(4.3)	(4.2)	(14.0)	(12.9)	(11.2)	
	4									
$g(y) = 1 - \exp\{-(y/30)^{-4}\}:$										
Maximum AOD	4.5	4.5	4.7	12.2	11.7	11.8	28.4	28.7	26.4	
	(1.3)	(1.5)	(1.2)	(3.5)	(3.4)	(4.0)	(7.1)	(7.6)	(7.4)	
Mean AOD	10.7	10.8	11.0	39.2	38.3	38.4 1	13.6	114.1	110.5	
	(2.1)	(1.7)	(1.9)	(6.4)	(5.3)	(6.0)	(18.1)	(17.7)	(15.1)	
Max reliable AOD	10.5	10.3	10.5	37.4	36.4	36.9 1	108.7	109.2	106.1	
	(1.9)	(1.6)	(1.7)	(5.2)	(4.4)	(4.9)(	(15.3)	(15.1)	(14.2)	

Table 2. Estimates of mean group size: sample mean (standard error in parentheses) of recorded group sizes within w of the line.

Mean group size	3			10			30		
Half-group spread	10m	25m	50m	10m	25m	50m	10m	25m	50m
$g(y) = 1 - \exp\{-(y/y)\}$	(20) <sup>-2</sup> }:								
			2.33						13.29 (0.82)
$g(y) = 1 - \exp\{-(y/30)^{-4}\}:$									
			2.52						

Fig. 1. Shown here are two datasets, both generated from a detection function with certain detection out to 40m, and rapidly declining detection probability at larger distances. When by chance there are more detections close to the line, visual inspection of the data leads to selection of a smaller cutpoint for the Kelker method; 20m in this example. When there are more detections close to 40m, the cutpoint is likely to be set at 40m. If we fix the cutpoint in advance, at either 20m for both analyses or 40m, we expect unbiased estimates of density, but if we use 20m for the first analysis and 40m for the second analysis, we overestimate density on average. Dashed lines: mean count with truncation at 40m. Dotted lines: mean count with truncation at 20m.

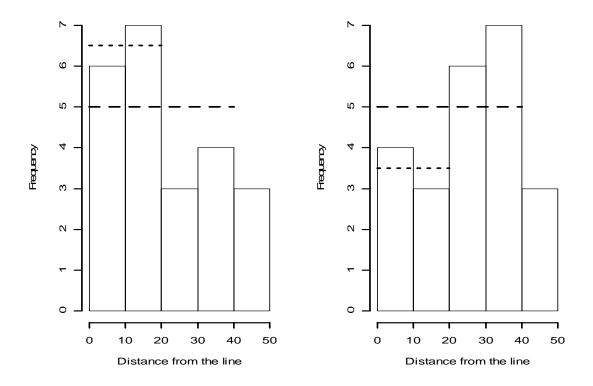


Fig. 2. The detection functions used in the simulation study. Note that these detection functions apply to each individual animal; the probability that at least one animal of a group will be detected will be larger than shown here – substantially so for large groups. The solid line is  $g(y) = 1 - \exp\{-(y/20)^{-2}\}$  and the dashed line  $g(y) = 1 - \exp\{-(y/30)^{-4}\}$ .

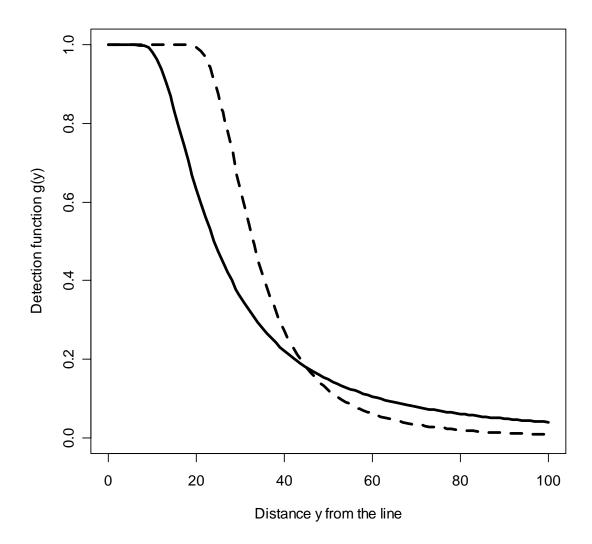


Fig. 3. Histogram of AODs simulated from the hazard-rate model of the detection process,  $g(y) = 1 - \exp\{-(y/20)^{-2}\}$ .

